

# **The U.S. left behind? Financial globalization and the rise of IPO activity around the world**

by

CRAIG DOIDGE, G. ANDREW KAROLYI, and RENÉ M. STULZ\*

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\* Doidge is at the Rotman School of Management, University of Toronto. Karolyi is at the Johnson Graduate School of Management, Cornell University. Stulz is at the Fisher College of Business, Ohio State University, ECGI, and NBER. Yeejin Jang, Rose Liao, and Taylor Nadauld provided excellent research assistance. We are grateful for comments from Steven Davidoff, Mathijs van Dijk, Crocker Liu, Edith Liu, Qing Ma, Maureen O'Hara, Elias Papaioannou, Raghu Rau, Stefano Rossi, and from seminar participants at Cambridge University, Cornell University, University of Florida, Harvard University, HEC Paris, Oxford University, Temple University, University of Amsterdam, University of Syracuse, and conference participants at the University of North Carolina Global Issues in Accounting Conference. Doidge thanks the Social Sciences and Humanities Research Council of Canada for financial support.

## **ABSTRACT**

During the past two decades, there has been a dramatic change in IPO activity around the world. Though vibrant IPO activity, attributed to better institutions and governance, used to be a competitive advantage of the U.S., it no longer is. IPO activity in the U.S. has fallen compared to the rest of the world and U.S. firms go public less than expected based on the economic importance of the U.S. Global IPOs, which are IPOs in which some of the proceeds are raised outside the firm's home country, play a critical role in the increase in IPO activity outside the U.S. The quality of a country's institutions is positively related to its domestic IPO activity and negatively related to its global IPO activity. However, home country institutions, viewed in earlier research as an explanatory factor for the historically high U.S. IPO rate, lose their explanatory power in explaining IPO activity as financial globalization increases. The evidence is consistent with the view that access to global markets and, more generally, financial globalization helps firms overcome the obstacles of poor institutions.

It is widely believed that a vibrant market for initial public offerings (IPOs) is an asset of the U.S. Black and Gilson (1998) and many others argue that the existence of such a market plays a critical role in facilitating entrepreneurship and venture capital in the U.S. economy. This view permeates corporate finance textbooks. For example, Megginson and Smart (2009) write: “Given its role in providing capital market access for entrepreneurial growth companies, the U.S. initial public offering market is widely considered a vital economic and financial asset.” The law and finance literature shows that IPO activity depends on country-level laws and governance institutions. For instance, it shows that IPO activity is higher in common law countries compared to countries with other legal origins. From this perspective, IPO activity has been vibrant in the U.S. because of better laws and better governance institutions. However, financial globalization makes it possible for firms to borrow institutions from other countries, so that we would expect that a firm’s home country institutions could become less important with financial globalization (see, among others, Coffee, 1999; Stulz, 1999; Doidge, Karolyi, and Stulz, 2007; and Stulz, 2009).

In this paper, we show that there has been a striking evolution over time in IPO activity across countries. We build a comprehensive sample of 29,361 IPOs from 89 countries constituting almost \$2.6 trillion (constant 2007 U.S. dollars) of capital raised over 1990 to 2007. Although the worldwide share of IPO activity by U.S. firms still ranks near the top during the 2000s, U.S. IPOs have not kept up with the economic importance of the U.S. In the 1990s, the yearly average of the number of U.S. IPOs comprised 27% of all IPOs in the world while the U.S. accounted for 27% of world Gross Domestic Product (GDP). Since 2000, the U.S. share of all IPOs has fallen to 12% whereas its share of worldwide GDP has averaged 30%. The average size of a typical IPO in the U.S. is larger than that in the rest of the world so that IPO proceeds may be a more relevant metric. Yet, in the last five years of our sample, IPO proceeds raised by U.S. firms drop to 16.2% of world IPO proceeds, despite the fact that the stock market capitalization of the U.S. relative to that of the world averages 41% during this period.

Some of the decrease in the importance of U.S. IPO activity compared to worldwide IPO activity may be due to lower IPO activity by U.S. firms, but much of it is explained by the considerable growth of

IPOs in other countries that occurs throughout the sample period. To a large extent, this growth is fueled by the emergence of global IPOs, which are those in which some or all of the shares are sold outside the home country of the firm going public. In 2007, proceeds raised in global IPOs accounted for 61% of total IPO proceeds, which is double the fraction raised in 1990. U.S. firms have never been active participants in the global IPO marketplace. This newer global IPO phenomenon is an important tool linked to the globalization of capital markets.

The most established literature that investigates cross-country differences in IPO rates is the law and finance literature (e.g., La Porta, Lopez-de-Silanes, Shleifer, and Vishny, hereafter LLSV, 1997, 1998). It focuses on how differences in countries' laws, governance, disclosure, and enforcement standards (hereafter "institutions") that protect minority shareholders can explain variation in IPO activity across countries. It predicts that IPO activity in countries other than the U.S. should increase relative to that in the U.S. if their institutions improve relative to those of the U.S. While there have been changes in laws and institutions in many countries, there is little evidence that these laws and institutions have become similar to those of the U.S. Alternatively, the institutions of the U.S. could have become worse, but there is no convincing evidence of such a development. A third possible explanation is that national institutions became less important as a determinant of domestic IPO rates because of financial globalization. Of course, lower IPO activity in the U.S. could also be explained by some factors that are not related to the quality of institutions but that are specific to the U.S.; such factors, however, cannot be readily accounted for in a cross-country framework.

The seminal law and finance papers study security issuance activity in a world with limited financial globalization. Hence, it is an open question as to how IPO activity is related to home country institutions in a world with dramatically reduced barriers to international investment and trade in financial services. Indeed, a newer literature, which includes Shleifer and Wolfenzon (hereafter SW, 2002), Doidge, Karolyi, and Stulz (hereafter DKS, 2007), and Stulz (2009) addresses the impact of financial globalization on IPO activity and suggests that home country institutions may have opposite effects on domestic compared to global IPOs. That is, global IPOs can be used to overcome the adverse effects of poor home country

institutions. In addition, we would expect that home country institutions should be less important in a more global world as firms can benefit from institutions and resources from other countries in their governance, even if they do not go public through a global IPO (Stulz, 1999). For instance, because of globalization, firms that go public in their own country can now use foreign auditors, law firms, and investment banks. The IPO literature emphasizes the importance of certification of the issuing firm (Ritter and Welch, 2002) and the use of foreign advisers and monitors can help certify the quality of the issuing firm in a more credible way than local advisers and monitors.

We use three different approaches to try to understand cross-country IPO activity around the world, its evolution over time, and the role of financial globalization. First, we conduct a comprehensive analysis of the role of country institutions for the cross-country variation of domestic IPOs. Second, our sample period allows us to evaluate the impact of increasing financial globalization on the role of institutions on IPO activity around the world. Finally, we investigate the relation between global IPO activity by firms from a country and the institutions of that country.

There are many theories of the decision to go public (see Ritter, 2003 for a review), but most of these theories ignore differences in institutions across countries. SW (2002) provide the archetypal model of how a country's laws and governance affect the benefits and costs of going public for the owners of firms and hence affect the likelihood that a firm will go public in a given country. In their model, the problem for public firms is that the controlling shareholder can extract private benefits at the expense of minority shareholders. However, minority shareholders buy shares at the IPO at fair value so that any expected private benefit consumption reduces IPO proceeds. At one extreme, laws and governance are so poor that any money provided by outside shareholders is consumed in private benefits. In such a situation, no IPO is possible. At the other extreme, if laws and their enforcement are so strong that no private benefits are consumed, firms that go public have high values and all entrepreneurs with positive NPV projects can go public. In reality, countries are between these extremes. Private benefits are lower in countries with good laws and good governance, so that in these countries the equity of firms is worth more and more firms benefit from going public.

Stulz (2009) adds an intermediate period to the SW model, where the entrepreneur has information that outside shareholders do not have about whether it makes sense to continue undertaking the firm's project. The entrepreneur benefits from continuing the project even when its NPV has become negative because he can extract private benefits upon completion of the project. Because of this problem, firms raise less capital at the IPO unless they can credibly commit to continuing the project only if it has a positive NPV. Stulz shows that a credible disclosure commitment can perform that role, but that laws must make it possible for outside shareholders to act if the news disclosed is negative about the project.

We test these predictions from the SW (2002) and Stulz (2009) models using several different country-level proxies for institutions based on laws, governance, disclosure, and enforcement standards. Countries with better institutions have more domestic IPO activity, measured as either the annual number of domestic IPOs scaled by the lagged number of domestic listed firms or as the annual proceeds raised in domestic IPOs scaled by lagged GDP. Firms can supplement weaker institutions in their home country through additional commitments to firm-level governance. DKS (2007) show that a firm's investment in mechanisms to improve corporate governance depends critically on the level of its home country's economic development, financial development, and openness. When economic and financial development is high, investment in governance is cheaper and more effective, so that a country's institutions are less important. We account for a country's per capita GDP, stock market capitalization to GDP ratio, and stock market turnover to measure the level of economic and financial development. Our results hold after controlling for these measures of economic and financial development.

The free flow of capital globally allows firms to raise funds publicly outside their country of domicile. By all measures, finance became much more global when one compares the 2000s before the global financial crisis of 2008-2009 to the 1980s. In the 1980s, many countries with viable stock markets were actually closed to capital flows (see Karolyi and Stulz, 2003). Very few of these countries were still closed or had substantial obstacles to capital flows in the 2000s. In the models of DKS (2007) and Stulz (2009), financial market globalization allows firms to borrow the institutions of foreign countries. Firms from countries with weaker institutions should benefit more from borrowing stronger institutions of

foreign countries. Therefore, we expect that such firms are more likely to go public with a global IPO and to raise more proceeds in foreign markets. Measuring annual global IPO activity by firms in a given country as either the ratio of the number of global IPOs to the total number of IPOs or as the ratio of global IPO proceeds to total IPO proceeds, we find strong support for the prediction that countries with weaker institutions have more global IPO activity. These findings are robust to controlling for other important determinants of IPO activity.

The importance of home country institutions for the extent of IPO activity can change over time with improved technology and with growing financial globalization. The models of SW (2002) and DKS (2007) predict that, with more capital market openness and globalization, the role of home country institutions for domestic IPO activity will diminish in importance. Specifically, DKS predict that the role of institutions for IPO activity, conditional on the level of financial and economic development, is lower when global markets are more accessible. To test this prediction, we construct a measure of financial globalization from the updated and extended dataset compiled by Lane and Milesi-Ferretti (LMF, 2007). Each year, we sum the U.S. dollar-denominated value of external assets and liabilities across countries and divide the sum by world GDP (also in U.S. dollars). Strikingly, the domestic IPO rate is strongly negatively related to financial globalization, while global IPOs increase with financial globalization. Perhaps even more importantly, the impact of national institutions on the domestic IPO rate falls strongly as financial globalization increases. For instance, using the anti-self-dealing index of Djankov, La Porta, Lopez-de-Silanes, and Shleifer (hereafter DLLS, 2008) as a measure of institutions, we find that a one standard deviation higher index score is associated with a 1.56% higher domestic IPO rate when the financial globalization measure is lower (such as during the 1990s when the ratio of external assets and liabilities to world GDP averaged 139%). However, a similar one standard-deviation higher index score is associated with only a 0.99% higher domestic IPO rate when this LMF globalization index is higher (such as during the 2000s when the ratio of external assets and liabilities to world GDP exceeded 252%). We also offer some evidence that home country institutions become a less important factor for the decision to

pursue a global IPO as a result of the increase in financial globalization, but these results are weaker than those for domestic IPO activity.

The literature on time-series variation in IPO activity focuses on changes in growth opportunities and market conditions. Ritter's (2003) survey points out that swings in the volume of IPOs are of considerable interest and that the volume seems to be "hypersensitive to changes in market conditions" (p. 293). Lowry (2003) addresses why IPO volume fluctuates so much and concludes that changes in aggregate capital demands of private firms and in investor optimism are the primary determinants. Pagano, Panetta, and Zingales (1998) find that for a sample of Italian IPOs the predominant reason firms go public is to rebalance their capital structure and to exploit mispricing, rather than to raise capital for financing investments. Loughran and Ritter (1995) also find support for the market-timing explanation for U.S. IPOs, while Henderson, Jegadeesh, and Weisbach (2006) find similar results internationally. To capture changing market conditions, we control for the country-level Tobin's  $q$  ratio (adjusted by the industry composition in that country) and a world domestic IPO factor.

Our tests and findings are new, but our effort is related to and adds to several recent papers. First, LLSV (1997), La Porta, Lopez-de-Silanes, and Shleifer (hereafter LLS, 2006), and Djankov, La Porta, Lopez-de-Silanes, and Shleifer (2008) show that legal protections for minority investors through rules and enforcement, as well as securities laws mandating disclosure and facilitating private enforcement, are associated with more IPO activity (on average, between 1996 and 2000). We use their legal protection and securities law measures, but we also capture important dynamics of IPO activity over time in conjunction with changing investment opportunities and with financial globalization. Further, we distinguish between domestic and global IPOs. Kim and Weisbach (2008) use a broad sample of 17,226 IPOs from 38 countries to examine firms' motives for going public. They find that IPO proceeds are predominantly used to finance investments and that there are few differences in the use of IPO proceeds between firms in common law and civil law countries. Caglio, Weiss Hanley, and Marietta-Westberg (2011) show that global IPOs account for a significant fraction of total IPO proceeds. They show that

characteristics of firms that choose a global offer are different from those that choose a domestic offer and that global IPOs originate from countries with lower bond and stock market development.

## **I. The IPO sample and country-level data.**

### *A. IPO data.*

We obtain IPO data from the Securities Data Company's (SDC) Global New Issues Database. For each IPO, this database provides information on the issuer, the issue date, total proceeds, the number and type of shares offered, the offer price, whether the issue is domestic only or contains an international tranche, and whether or not a tranche is offered to public or private investors. We begin by downloading all transactions in SDC where the "original IPO" flag is set to "yes." Because SDC has very limited coverage for non-U.S. IPOs prior to 1990, our sample begins in January 1990. The sample ends in December 2007. The initial count is 38,724 observations. We eliminate transactions with a single domestic tranche that SDC flags as a private placement (57 observations). There are 526 cases in which there is more than one transaction reported in SDC for the same firm within a narrow window of time. Many of these are global IPOs where the domestic and international tranches have different issue dates, usually within a few days of each other. We drop 235 observations with a gap of 30 days or more between issue dates. Following Kim and Weisbach (2008), we remove 48 transactions that do not contain any information on proceeds raised. The data for some IPOs is recorded over multiple lines in SDC, even if there is only one tranche in the offering. We consolidate these issues into one line of record and drop 1,347 observations. Some foreign IPOs (those with an international tranche but no domestic tranche) and all global IPOs (those with both domestic and international tranches) are also recorded over multiple lines in SDC. We consolidate that information into one line and drop the 3,638 duplicate records. We also drop 93 transactions that do not have SIC codes, leaving us with 33,306 observations, each of which represents a unique IPO.

To construct our final sample, we exclude an additional 3,945 IPOs. We drop 3,856 IPOs by REITs and investment funds, 44 IPOs where the country of origin has no data (more details below, but they

include tax havens like the British Virgin Islands, Guernsey, or Cayman Islands), and 45 IPOs from 16 countries for which there were no domestic IPOs (only global IPOs) during the 18-year sample period.<sup>1,2</sup> The final sample contains 29,361 IPOs from 89 different countries of which 24,122 are purely domestic and 5,239 are foreign only and global offers.

While the sample just described is the one used in the tables we report, we repeat our tests for a sample that excludes all privatizations and for a sample that counts Chinese IPOs that list in Hong Kong as domestic IPOs even though they are subject to different laws and regulations as a result of the Hong Kong listing. Since our regressions are at the country-year level, these alternative samples do not reduce materially the number of observations in our regressions. The results for these alternative samples are consistent with the results we report.

We perform two experiments to lend assurance that our SDC sample is a reliable representation of IPO activity around the world. In one experiment, we collect data from the World Federation of Stock Exchanges (WFE). Each year, the WFE surveys their member, affiliate, and correspondent exchanges on a wide range of statistics, including what they call “investment flows”. This includes new companies that list and the new capital that they raise via shares. The WFE provides a list of definitions and calculation methods that the exchanges must follow to increase the comparability of the information across exchanges (see [www.world-exchanges.org/statistics](http://www.world-exchanges.org/statistics)). We obtained this data for 2001 through 2005 for each country and compared it with our IPO counts and proceeds totals from SDC. In general, SDC and WFE have comparable coverage and reporting for many countries. North American and U.K. totals are close, as are those for most Asian countries, such as Singapore, China, and Hong Kong. There are patterns of SDC under-reporting of counts and proceeds in Australia, India, Italy, Spain, Sweden, and Switzerland. Of course, there may be good reasons for these differences, such as the inclusion of investment funds and REITs in the WFE samples. Additional complications arise in the WFE’s data with the consolidation of

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<sup>1</sup> The following SIC codes were used to screen out REITs and investment funds: 6722, 6723, 6726, 6798, and 6799.

<sup>2</sup> Countries with no domestic IPOs are: Angola, Barbados, Cambodia, Dominican Republic, Faroe Islands, Georgia, Ghana, Iceland, Kazakhstan, Lebanon, Macau, Malta, Netherlands Antilles, Slovenia, Ukraine, and Uruguay.

exchanges, such as for NYSE Euronext and OMX Nasdaq, as they centralize their reporting relationship with the WFE. The WFE does not report on the composition of IPOs by type.

We also collect IPO data from Bloomberg and from the home-market exchanges for four randomly chosen countries (Brazil, Canada, Germany, and Malaysia) from the early 1980s through 2007. In each case, we obtain information on domestic and global IPOs, but only by counts not proceeds. For Germany from 1997 to the present, the SDC counts are almost identical to those reported by the Deutsche Börse on their website. Those for Bloomberg are higher (about 50% discrepancy on average); they report more than double the count in 2005-2006 relative to the Deutsche Börse and SDC.<sup>3</sup> For Malaysia, the IPO counts in Bloomberg are very similar to those from the Bursa Malaysia website and SDC (less than 5% discrepancy, on average). The WFE reports a much higher count for Brazil's Bovespa than Bloomberg and SDC, though the latter two are similar (about 30% discrepancy, on average). Finally, for Canada, the Bloomberg counts are on average 40% lower than for SDC which are, in turn, about 20% higher than those reported to the WFE. The biggest count discrepancies occur during 2000 and 2001.

#### *B. Country level data.*

In our regressions, the dependent variable is a measure of the rate of IPO activity. For each country and each year, we compute the number of IPOs ("IPO counts") as well as the total proceeds raised in IPOs ("IPO proceeds"). To compute the IPO counts and proceeds, we distinguish between domestic IPOs and global IPOs (foreign IPOs are included with global IPOs). To benchmark IPO activity across countries that differ in size, we scale the IPO counts by the lagged number of publicly-listed domestic companies the country of domicile and the proceeds raised each year by lagged home country GDP. These data are obtained from the World Bank's World Development Indicators (WDI) Database. Listed domestic companies include domestically incorporated companies listed on the country's stock exchanges at the end of the year and do not include investment companies, mutual funds, REITs or other collective

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<sup>3</sup> Each of the 81 German IPOs listed in Bloomberg in 2005 were manually checked and several firms (e.g., Bertelsmann, IC Immobilien, Marenave Schifffahrts, and Qsil) were not on the Deutsche Börse website. These firms had announced plans to do an IPO, but subsequently announced that they would defer the IPO due to restructuring or other reasons. Bloomberg appears to rely on corporate news releases and prospectuses.

investment vehicles. GDP is reported in U.S. dollars converted from domestic currencies using the end-of-year official exchange rate for that country.

An important set of data in our work are country-specific institutional variables related to the quality of investor legal protections and securities laws related to disclosure requirements and enforcement standards. From LLSV (1998), it is well-known that common law countries have better institutions. We therefore use the common law dummy introduced by LLSV (1998) and extended in DLLS (2008). It equals one if the origin of commercial law in a country is English common law, and zero, otherwise (“Common law”). A popular index of legal protections for minority investors is the anti-director rights index (“Anti-director”) of LLSV (1998) and updated and revised by DLLS (2008) based on laws that apply to firms in 2003. The index is formed on a six-point scale based on a set of variables meant to capture the stance of corporate law towards shareholder protection.<sup>4</sup> DLLS (2008) build an index of anti-self-dealing (“Anti-self-dealing”) to address the ways in which the law deals with corporate self-dealing in a more theoretically grounded way. It is assembled by means of a 2003 survey of Lex Mundi law firms in 72 countries and includes components related to ex ante private control of self-dealing, such as disclosures that counterparties in a transaction must make before approval is granted by disinterested shareholders as well as similar ex post disclosures (such as access to evidence) for independent reviews of transactions after completion toward possible rescission or follow-on suits.

LLS (2006) show that securities laws that mandate prospectus disclosure and prospectus liability benefit stock market development, including the breadth, size, and liquidity of the market. They devise measures based on a survey of attorneys in 49 countries in 2000. These measures are especially useful for our study as they relate closely to the security issuance process in IPOs. They build a disclosure requirements index (“Disclosure”) with components related to requirements for prospectuses, and for providing information on compensation of directors and key officers, the issuer’s ownership structure, related-party transactions with directors, officers or large blockholders, and the presence of contracts

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<sup>4</sup> Spamann (2010) also re-codes the original anti-director rights index used in LLSV (1998). He does not report data for all countries in his sample, so we do not employ this alternative index.

outside the ordinary course of business. The liability standard index (“Burden of proof”) comprises measures of four liability standards in cases against issuers and directors, distributors, and accountants. The index of public enforcement (“Public enforcement”) is based on five broad aspects of public enforcement: the basic characteristics of the supervisory body for securities markets, the scope of its powers to regulate markets, its investigative powers, its power to issue noncriminal sanctions for violations of securities laws against issuers, distributors and accountants, and whether, to whom, and when criminal sanctions for violations of securities laws apply. Finally, LLS build an all-encompassing investor protection index (“Investor protection”) which is the first principal component of the burden of proof, disclosure, and the anti-director rights index from LLSV (1998).

We also include a measure of the rule of law (“Rule of law”) from the World Bank’s World Governance Indicators database<sup>5</sup> and political risk (“Political risk”) from the International Country Risk Guide (ICRG) database built by The PRS Group, Inc. In contrast to the LLSV and DLLS variables, these variables are measured every year. The former captures perceptions of the extent to which agents have confidence in and abide by rules related to contract enforcement, property rights, the police and the courts as well as likelihood of crime and violence. It is based on a survey of public and private sector experts and is available for over 200 countries since 1996, including annually from 2002. The political risk variable from ICRG includes components related to government stability, socioeconomic conditions, internal and external conflicts, corruption, law and order, democratic accountability and bureaucratic quality. It is scored on a 100-point scale and is available annually from 1990. This political risk variable serves as a proxy for the quality of country governance. To the extent that countries with high political risk are countries where public firms are more at risk of predation from the state (Stulz, 2005), we would expect IPO activity to be weaker in such countries.

A key mechanism through which poor institutions limit IPO activity in the literature we have discussed is that they require more co-investment by insiders at the IPO. Consequently, we would expect fewer IPOs in countries where insider ownership is optimally higher. We use a measure of insider

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<sup>5</sup> For details on methodology and analytical issues, see Kaufmann, Kraay, and Mastruzzi (2010).

ownership (“Ownership”) from LLSV (1998), which is computed as the average percentage of shares owned by the top three shareholders in the 10 largest, nonfinancial, private domestic firms in a country.

For our analysis, we require a measure of the extent of financial globalization. The literature uses a number of measures for *de jure* openness. While these measures exhibit cross-sectional variation, they generally have only limited time-series variation during our sample period. Moreover, they do not capture the extent to which financial globalization is actually taking place (*de facto* openness), which is central to our analysis. If residents from one country can invest freely in another country, they may or may not do so. If they do not do so, it is not clear why financial globalization should lead to more global IPOs since there would be no appetite for IPOs by foreign companies. A proper measure for our purpose should then focus on the extent of gross external claims, including assets and liabilities.<sup>6</sup> We therefore use the measure of financial globalization from the updated and extended dataset compiled by Lane and Milesi-Ferretti (2007). For each country, it is computed as the sum of a country’s external assets and liabilities divided by the GDP (“Country financial globalization”). By computing the sum across each country’s external assets and liabilities and dividing by world GDP each year, we also have an annual measure of world financial globalization (“World financial globalization”). This measure equals 116% of GDP in 1990. It increases slowly until 1997, when it equals 149%. From 1997 to 2000 it increases sharply to 204%, after which it is relatively stable for a few years. It then increases sharply again and reaches 327% of GDP in 2007, the last year of the sample period. The regressions we report use the world financial globalization measure. We discuss additional results using a different measure that includes only equity assets and liabilities (excluding debt claims and foreign currency reserves) as well as with the country financial globalization measure.

In our regressions, we also include measures of development. To measure the level of economic development in the country, we use the log of GDP per capita (“Log(GDP / capita)”). This variable is obtained from the WDI Database. For measures of financial market development, we use the 2008 update

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<sup>6</sup> A measure of net cross-country claims is not appropriate since a country with zero net claims on foreign residents could be a country with large foreign equity asset holdings financed with foreign debt liabilities. See Borio and Disiyata (2011) for a discussion of the limitations from using measures of external net claims.

of the Financial Development and Structure database, originally used in Beck, Demirgüç-Kunt, and Levine (2000). We collect data for the stock market turnover ratio (“Market turnover,” the ratio of the value of total shares traded to average real market capitalization) and stock market capitalization as a percentage of GDP (“Market cap / GDP,” the ratio of the market capitalization of listed shares to GDP).

To control for local market conditions as a factor in the going-public decision, we compute a country-level measure of Tobin’s  $q$  each year. At the firm level,  $q$  is computed from data obtained from Thomson Reuter’s Worldscope database as follows: the numerator is total assets less the book value of equity plus the market value of equity. For the denominator, we use the book value of total assets. All variables are in local currency. Using the Fama-French 17 industry classification scheme, we compute the median  $q$  and relative market value for each industry annually. The country-level measure of  $q$  (“Country  $q$ ”) is the market value weighted average of the median industry  $qs$ . This measure is constructed analogously to the local growth opportunities (LGO) measure based on P/E ratios used in Bekaert, Harvey, Lundblad, and Siegel (2007).

Finally, we construct a measure to control for unobservable macroeconomic and capital market factors that influence IPO activity around the world (“World domestic IPO rate”). It is measured either in terms of IPO counts per number of listed firms or in terms of IPO proceeds per GDP. Each year, total world IPO counts (proceeds) are summed across countries and scaled by the lagged total number of listed firms worldwide (lagged world GDP). To compute the world IPO rate for a given country, the IPO activity and its associated scale factor (either number of listed firms or GDP) for that country is excluded.

Summary statistics, including means, standard deviations, and correlations among these measures are provided in Appendix A.

## **II. The rise of IPOs abroad and the fall of IPOs in the U.S.**

Table 1, Panel a presents the total number of IPOs and breaks it down by domestic IPOs and global IPOs by year. Annual IPO counts increase from less than 1,000 in the early 1990s to a peak of 3,100 in 1996. They decrease after 1996 before reaching another peak of 2,117 in 2000. The counts fall below

1,000 for three years after 2000 before increasing again steadily to reach 1,850 in 2007. Panel a also shows that the rise and fall in annual counts until 2003 occurs for both domestic and global IPOs. The surge in overall counts after 2003 is much more dramatic for global offers. For domestic IPOs, 2007 does not exceed the earlier peaks, while the count for global IPOs in 2007 is higher than in any other year in the sample period. Generally, there is more year-to-year variation in global IPOs. From the peak in 2000 to the trough in 2003, global IPOs fall by 84%; in contrast, domestic IPOs fall by only 44%.

The results for IPO proceeds are presented in Table 1, Panel b. We obtain proceeds in U.S. dollars from SDC and convert them into constant 2007 values using U.S. inflation rates from the World Bank's WDI database. This panel shows that total annual IPO proceeds rise during the 1990s to reach a peak of \$240 billion in 2000. Annual proceeds decline to \$59 billion by 2003 and then rise again, reaching a peak of almost \$280 billion in 2007. Domestic IPO proceeds are less volatile over the period, so that changes in annual proceeds from global IPOs are the more important factor in the steady rise of total IPO proceeds during the 1990s and especially in the rapid expansion after 2003, reaching \$189 billion in 2007. Indeed, total proceeds raised in global IPOs account for almost 68% of all IPO proceeds in 2007. Global IPOs include a domestic tranche and international tranches. The last column of Panel b shows the proceeds that are raised in the international tranches of global IPOs. As a percentage of total proceeds raised in global IPOs, proceeds raised in international tranches have increased over the 2000s, reaching a peak of 90% in 2007.

There are some important differences in the evolution of counts and proceeds in both domestic and global IPOs. In the 1990s, there is a dramatic increase in counts that is driven by an increase in domestic IPOs. The number of domestic IPOs peaks in 1996 and does not come close to that peak again in subsequent years. In fact, after 2000, the count never exceeds even half the peak reached in 1996. In contrast, however, domestic IPO proceeds are higher in the mid-1990s, but the proceeds raised in 2006 dwarf those of earlier years. There is a steady increase in the number of global IPOs until 2000. The count then drops, but increases again after 2003 and peaks in 2007. Proceeds that are raised in global IPOs increase throughout the 1990s to reach a peak in 2000. They collapse to a trough in 2003, and then

increase sharply to reach the peak of 2000 again in 2007, although the percentage of global IPO proceeds that are raised in international tranches is much higher in 2007 compared to 2000.

The cross-country pattern in annual IPO counts is exhibited in Panel a of Table 2. Developed countries with the largest economies and capital markets in the world, such as the U.S. (6,126 IPOs), Japan (2,234), Canada (2,225), U.K. (1,650), Australia (1,558), and Hong Kong (822), have the highest overall counts, but a number of emerging countries such as India (4,867), China (1,764), Taiwan (822), and South Korea (779), have high counts as well. Fifteen countries are in the top 25 both for counts and proceeds. Panel b shows that the U.S. total of \$648 billion constitutes almost 25% of the total worldwide IPO proceeds of \$2.55 trillion. The other major markets include China (\$254 billion, 10%), Japan (\$204 billion, 9%), U.K. (\$196 billion, 8%), and are followed by France (5%), Germany (4%), Canada (3%), and Italy (3%). However, some countries that are in the top 25 for counts are not in the top 25 for proceeds (Greece, Israel, Malaysia, Pakistan, and Poland), whereas some countries are in the top 25 for proceeds but not for counts (Bermuda, Mexico, Russia, Spain, and Switzerland).

The country-by-country averages hide dramatic changes in the frequency of IPOs across countries. In addition to showing the counts for all IPOs across the world, Figure 1 shows the counts for the U.S., the U.K., and China (Panel a). The U.S. dominates the U.K. and China in counts until 2001 when the counts are roughly the same and stay that way to the end of the sample period. The U.S. counts peaked in 1996, but the number of U.S. IPOs has been small compared to that peak since 2001. If at one point the U.S. was the “land of the IPO,” it is not in the 2000s if one focuses on IPO counts. One way to see this is that the U.S. share of total IPO counts exceeds 20% in each of the first ten years of the sample except for 1994 when it is 19%. It never exceeds 13% after 2001. In the 1990s, the U.S. share of total IPO counts towers over the share of the U.K. and China, as each country’s share is below 10% each year in the 1990s. Though the shares of these countries increase in the 2000s, the U.S. share stops towering over their shares, mostly because the U.S. share is so much lower in the 2000s. Japan, Australia, Taiwan, and Korea all experienced substantial increases in counts as well that we do not show on the figure.

There has also been a dramatic shift in the composition of IPO proceeds over the past two decades. The U.S. share of total IPO proceeds has declined from about 30% in the 1990s to only 21% in the 2000s (through 2007, at least). Japan and the U.K. have also experienced a decline from 10% to 6% and from 9% to 6%, respectively. Among the other large markets, no major shift in market share arises (e.g., Canada, France, Italy, and Germany), except for China which more than doubles from a 6% to a 14% share (\$182 billion out of the \$1.29 trillion). In 2006 and 2007, China's total IPO proceeds actually exceed those of the U.S. (see Panel b). Because of the dramatic changes in China's IPO market, we perform supplementary tests throughout our study for a sample that excludes Chinese IPOs.

The decreased importance of U.S. IPOs occurs at different times for counts versus proceeds. The share of U.S. counts in world IPO counts decreases sharply until 1994. After that year, it increases until 1999 and then collapses starting in 2000. It stays steadily low in the 2000s. In contrast, U.S. IPO proceeds mostly parallel the world's proceeds until 2003, at which point the world IPO proceeds take off and the U.S. IPO proceeds do not. One useful way to compare patterns in counts and proceeds for the U.S. and the world is to benchmark them relative to the number of listed firms and GDP, respectively. Figure 2 summarizes this evidence. It reports the ratio of IPO counts to the number of domestic listed firms (Panel a) and the ratio of IPO proceeds to GDP (Panel b) for both the U.S. and for the world. It also reports the difference between the U.S. and the world. The U.S. tends to have higher values than the world, but not consistently, until 1999. After 2003, the ratios for the world increase while the ratios for the U.S. stagnate. If the importance of IPOs in the U.S. were measured relative to the economic importance of the U.S. and if this ratio were benchmarked against the equivalent one for the world, we might expect these ratios to stay constant over time. Instead, we see that the ratios fall over time, so that the importance of IPOs in the U.S. relative to the world has not kept up with the economic importance of the U.S.

Figure 3 performs a different benchmark analysis for the U.S. by computing its share of IPO counts and proceeds relative to that of the world over time. The statistics are reported separately for all IPOs, domestic IPOs only, global IPOs only, and also for the global component of global IPOs for the analysis based on proceeds. We observe a steady decline in the U.S. share of IPO counts (Panel a) regardless of

the type of IPO. For the U.S. share of the world's IPO proceeds (Panel b), we see that the decline arises primarily from the share of global IPO proceeds from around 10% in the 1990s to a negligible fraction after 2001. In fact, the U.S. holds a steady fraction of domestic IPO proceeds around 35% over these two decades.

The evidence in this section shows that IPOs in the rest of the world have become much more important and IPOs in the U.S. have become less important. One possible explanation could be that country affiliations have simply become less important because firms wanting to pursue IPOs have found ways to avoid being hindered by national institutional obstacles. For instance, firms can use global markets to go public and avoid the constraints of their home country. A second possibility is that there is a “catching-up” effect. Foreign countries had lower rates of IPO activity in the past and they are simply catching up in the 2000s for reasons having little to do with the quality of national institutions. There is, after all, a vast literature that focuses on the role of investor sentiment and on growth opportunities for IPOs (Loughran and Ritter, 1995; Ritter, 2003), which may have influenced why IPO activity in the U.S. was more robust in the 1990s. A final possibility is that market conditions were relatively more attractive in the U.S. in the 1990s and became relatively more attractive in foreign countries in the 2000s. In the next section, we explore the importance of country characteristics as determinants of IPO activity. In Section IV, we address the role of global IPOs.

### **III. Do national institutions and market conditions matter for IPO activity around the world?**

In order to assess the importance of national institutions, like corporate laws, securities laws, their enforcement, and measures of political risk, as well as market conditions, such as equity valuations in a country, we need to benchmark IPO activity in terms of both counts and proceeds relative to the extent of potential activity. The literature has employed different approaches to gauge this potential activity. Previous work on IPO activity in a country in terms of counts has been benchmarked relative to the population in a given country (LLSV, 1997) and the number of listed companies on the major exchanges (DLLS, 2008). We choose the latter. IPO activity in terms of the proceeds of equity issued by newly listed

firms in a country has been benchmarked relative to its GDP by LLS (2006) and DLLS (2008) and relative to the total assets of the firms involved in raising capital (Kim and Weisbach, 2008). We choose the former.

Table 3 presents summary statistics by year for domestic IPO counts as a fraction of the previous year's number of domestic listed firms (Panel a) and for domestic IPO proceeds as a fraction of the previous year's GDP in millions of U.S. dollars (Panel b). Domestic IPO proceeds include proceeds from domestic IPOs only. We multiply both ratios by 100 and winsorize them at the 1% and 99% thresholds. We restrict the analysis to the countries used in Tables 1 and 2 that have sufficient data on changing market conditions that we include in our regressions below.<sup>7</sup> IPO activity by counts ranges from a low of 0.43% of listed firms in 1990 to as high as 5.38% in 1994. These are means across countries and it should be noted that there is significant dispersion in activity across countries by year and, moreover, that the number of countries with non-zero IPO counts changes over time. Fewer countries have no IPOs when IPO markets are hot around the world than when they are cold. Specifically, the number of countries that have no IPOs in a year is negatively correlated with the worldwide average of IPOs per number of listed firms.<sup>8</sup> This evidence suggests that world market conditions play an important role in IPO activity at the country level. We provide more such evidence in regressions that analyze the country-level IPO rate.

In Panel b, domestic IPO proceeds as a fraction of GDP ranges from as low as 0.08% in 1990 to as high as 0.26% in 1994. The time-variation in IPO proceeds across years follows closely the pattern in counts per number of listed firms, but not perfectly. This fact implies that there are interesting differences in the offering sizes of IPOs across years, part of which stem from the types of firms that go public and part of which stem from the countries of domicile that dominate IPO activity in those years. It is useful to

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<sup>7</sup> An important control variable in our subsequent analysis is country  $q$  for which we require that a country have this data available for at least one year to qualify for the summary statistics in Table 3. This restriction eliminates 35 countries (which together constitute only 300 IPOs out of our total sample of 24,122) leaving 54 countries. In supplementary tests, we expanded our analysis to include the country-years for which country  $q$  is not available. The overall explanatory power of our regressions decline, but key inferences associated with our control variables and national institutions proxies remain.

<sup>8</sup> We tested whether this relationship is statistically reliable using Tobit regressions of the percentage of countries with non-zero counts or proceeds on the mean IPO rate across countries by count or proceeds, respectively. There is a statistically significant negative coefficient for the proceeds relationship (coefficient of -1.30,  $t$ -statistic of -2.66) and a negative, but insignificant, coefficient for the same by counts.

point out that the range of this fraction is limited by the fact that proceeds (in billions of U.S. dollars) are typically small relative to the GDP of a country (also, in billions of U.S. dollars) even if scaled by 100. In particular, the maximum fraction of IPO proceeds never exceeds 2% in any year. Again, there is a large fraction of countries that are counted in these means by year for which there are no IPO proceeds.

*A. The role of changing investment opportunities and market conditions.*

Everything else being equal, we would expect to see more IPOs in countries with better growth opportunities, with more economic development, and with higher financial development. We include  $\log(\text{GDP} / \text{capita})$ , market cap / GDP, and market turnover as measures of the level of economic and financial market development, and lagged country  $q$  as a measure of corporate valuations in the country. Further, the world domestic IPO rate is included to control for unobservable macroeconomic and capital market factors that influence IPO activity around the world. Finally, we include world financial globalization as we expect that greater financial globalization makes it easier for firms to use global IPOs to go public and thereby rely less on raising capital using a domestic IPO.<sup>9</sup> Each of these variables is lagged by one year. Lagging these variables is especially important for market capitalization and turnover since these variables would be directly affected by IPO activity. We estimate this specification as a panel regression using ordinary least squares allowing the standard errors to be clustered by country.

Table 4 presents the estimates for regressions that project measures of domestic IPO activity on domestic economic and financial variables as well as domestic IPO activity outside the country. The specifications in Panel a use domestic IPO counts normalized by the lagged number of listed firms as the measure of IPO activity and the specifications in Panel b use domestic IPO proceeds relative to lagged GDP.

In Panel a for domestic IPO counts, the first important finding is that the world domestic IPO rate is a reliably positive and economically important factor. The coefficient of 0.302 implies that a one standard deviation increase in the world IPO rate (2.24%) is associated with a 0.68% increase in the IPO rate for a

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<sup>9</sup> The world domestic IPO rate has little within variation across countries in a given year and world financial globalization has none. Therefore, we cannot include year dummies in the regressions.

given country, or about 15.4% of its standard deviation. This is a sizeable source of common variation in IPO activity around the world. We performed several robustness checks to confirm the reliability of this result. Another statistically reliable coefficient is that for country  $q$ . The economic magnitude of country  $q$  is somewhat larger than that of the world IPO rate. Its coefficient of 2.703 implies that a one standard deviation increase in  $q$  (0.33%) is associated with a 0.89% increase in IPO activity rate, which constitutes 20% of its standard deviation. The coefficients on market cap / GDP and market turnover are positive as expected. The coefficient on market turnover is significant at the 5% level while the coefficient on market cap / GDP is significant only at the 10% level. Log(GDP / capita) has a negative coefficient but it is not significant. Finally, world financial globalization has a negative coefficient that is statistically significant and economically important. A one-standard deviation increase in the world financial globalization measure (0.635%) is associated with a 0.41% decrease in the domestic IPO rate, or about 9.5% of its standard deviation. The explanatory power of this specification is reasonable with an adjusted  $R^2$  of 13.6%.<sup>10</sup> When we use a measure of world financial globalization based on the sum of equity assets and liabilities across countries as a fraction of world GDP (thus excluding debt claims and accumulated foreign currency reserves) our findings (not reported) are very similar to those shown in Panel a. As discussed earlier, we also use a country-specific measure of financial globalization from Lane and Milesi-Ferretti (2007). In unreported results, we find that its coefficient is also negative and significant, but only at the 10% level and its economic magnitude is smaller. The concern with the country-specific measure is that, even though it is lagged, it is likely to be higher for countries that have had global IPOs. In contrast, there is little risk of such a mechanical relation with the world measure.

The first column in Panel b presents the results of the same base regression for domestic IPO proceeds normalized by lagged GDP. The results are similar to those for counts. The world domestic IPO rate (excluding the country of interest) is reliably positive. Its coefficient of 0.634 implies that a one standard deviation increase in the world IPO activity rate is associated with an increase of 0.036%, which

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<sup>10</sup> We performed some residual diagnostics for the base specifications in Panels a and b of Table 4. There is evidence of significant non-normality of the OLS residuals. Later, we report on supplementary tests dealing with concerns about censored samples using Tobit regression analysis.

constitutes 12.5% of its standard deviation.<sup>11</sup> In this specification, the coefficient on country  $q$  is also positive, statistically reliable, and has a larger economic magnitude corresponding to about 18% of the standard deviation of the IPO activity rate in proceeds. The coefficient on market turnover is positive and significant at the 10% level, while that for market cap / GDP is significant at the 1% level. Finally, world financial globalization has a statistically-significant negative coefficient that implies an economically large relation (11.9% of the standard deviation of domestic IPO proceeds). The explanatory power of this base specification is again reasonable with an adjusted  $R^2$  of 12.8%.

In supplementary, unreported regressions, we estimated specifications with annual U.S. dollar-denominated national index returns in excess of the Datastream World Index instead of country  $q$ . The coefficient is reliably positive yielding similar explanatory power. We also estimated specifications with and without  $\log(\text{GDP} / \text{capita})$  without much consequence. Further, we found that expanding the panel to include countries with no IPO activity during the 18-year horizon leads to similar conclusions. The next question we pursue is whether measures of the quality of national institutions can explain how domestic IPO activity differs across countries beyond these factors that relate to market conditions and the level of capital market and economic development.

*B. National laws and governance institutions and domestic IPO activity.*

In both Panels a and b of Table 4, the remaining specifications for the panel regressions of domestic IPO activity include one of a variety of country-level measures of laws and governance institutions. We include these variables one at a time, but have also estimated in unreported specifications various combinations of them. We should also state from the outset that our analysis below for domestic IPO proceeds as a fraction of GDP is robust to including domestic proceeds from global IPOs in the numerator.

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<sup>11</sup> We conducted a principal components analysis (PCA) of the first several principal components of IPO activity rates (for IPO proceeds as a fraction of GDP). The proportion of total variation across countries and years explained by the first PC was 9.61%, which is very close to the economic magnitudes we uncovered above. The second PC explained another 6.65%, the third PC, another 5.03% and the first five PCs together cumulatively explained 31.8%. These proportions were stable across different subperiods.

In general, we find that the better is the quality of the national institutions in a country the higher is the level of domestic IPO activity. In the second column of each panel, the addition of the common law dummy variable adds explanatory power to the base model. The coefficient is positive and significant at the 10% level with a value of 1.55 for domestic IPO counts as a fraction of the number of listed firms and positive and significant at the 5% level with a value of 0.11 for domestic IPO proceeds as a fraction of GDP. LLSV (1997) find similar results using counts of IPOs per millions in population for a two-year period, 1995-1996. In contrast to their approach, we use a panel regression and account for country and world market conditions. The positive coefficient on anti-director rights is insignificant for both the IPO count and proceeds regressions and the adjusted  $R^2$  is virtually unchanged with its inclusion. In LLSV (1997, Table VI), the coefficient on the original anti-director rights index is reliably positive for their IPO counts regressions, but in LLS (2006, Table III), it is similarly insignificant when they measure IPO proceeds relative to GDP (for 1996 to 2000), like we do in Panel b. The anti-self-dealing index from DLLS (2008) is positive and significant at the 1% level in both panels. For IPO counts, the coefficient of 5.602 implies that a one standard deviation higher score in anti-self-dealing (say, from that of Switzerland to Canada) is associated with an increase of 1.34%, or 31% of its standard deviation. The economic magnitude of this variable is similar for IPO proceeds. The adjusted  $R^2$  increases substantially to 21.19% (from 13.6% in the base specification) in Panel a and to 19.7% (from 12.8%) in Panel b.<sup>12</sup>

The next regressions include the securities laws measures developed in LLS (2006). We include the indexes for disclosure, burden of proof, public enforcement, and investor protection. LLS (2006) show these variables separately and together to be statistically and economically important for explaining IPO activity measured by proceeds for the high IPO activity period of 1996-2000 (their Tables III and V). We similarly find that the coefficients on these variables are important for both IPO counts and proceeds. For example, disclosure has a coefficient of 7.177 (robust  $t$ -statistic of 4.84) in Panel a for IPO counts; this implies that a one standard deviation higher score (say, from that of Turkey to Spain) is associated with a

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<sup>12</sup> In Table 6 of DLLS (2008), the anti-self-dealing variable has a reliably positive coefficient of 4.14 for their specification on IPO proceeds relative to GDP though with fewer control variables for the level of economic and financial development. They discuss the large economic significance of this variable (p. 449).

1.50% higher rate of IPOs to the number of listed companies, which represents about 34% of its standard deviation. The adjusted  $R^2$  for this specification in Panel a is 23.0% (from 13.6% in the base specification). The results for the public enforcement, burden of proof, and investor protection indexes are also statistically significant and large economically. That the results for the disclosure index are particularly strong is important support for the key prediction in Stulz (2009), where credible disclosure commitment *ex ante* and *ex post* by means of strong securities laws is critical for the entrepreneur to maximize offering proceeds.

We obtain similar, but distinctly weaker, results for several of the other national institutions variables we consider. Stulz (2005) argues that an entrepreneur has an incentive to not go public in countries where the state is predatory since it is easier for the state to prey on public companies as more information is available about them. We find no evidence to support this prediction in terms of IPO counts or proceeds. We also examine a measure of the rule of law. There is evidence of a positive relationship, but only for IPO counts (coefficient of 1.15 with *t*-statistic of 2.16). Finally, we evaluate a measure of ownership concentration. Such a measure is not an institution variable, but an outcome variable. A number of the theoretical models we work from associate the lack of ownership concentration with stronger laws and governance institutions. SW (2002) predict lower concentration of ownership in countries with better investor protections and Stulz (2005) associates this outcome with a less predatory government. In LLSV (1998), LLS (2006), and DLLS (2008), lower ownership concentration (computed as the stake of the top three shareholders in the largest 10 firms in a country) is associated with common law origins and higher scores on the anti-director rights, anti-self-dealing, disclosure, burden of proof, and public enforcement indexes. However, there could be other reasons why ownership is concentrated in a country. Whatever the reason, the more concentrated is ownership the less there is to gain from an IPO since the float is likely to be smaller. We find that ownership is reliably negatively related to domestic IPO activity for counts, but not for proceeds.

There are several potential concerns with the regressions of Table 4. First, the sample we use excludes a number of countries because of lack of data. We estimate (but do not tabulate) the regressions

using the largest sample possible for each institution variable. In these regressions, the sample increases by more than 50% for some variables (e.g., the anti-director index, the rule of law index), but does not increase meaningfully for others (e.g., public enforcement, ownership). Nevertheless, when we use these extended samples, our results are consistent with the results we report. Second, institutions could change because of a demand for IPOs, so that the quality of institutions in a country could be endogenous to IPO activity that takes place there. This concern is not plausible with the common law dummy. Yet, we find evidence of an association of common law with IPO frequency. The concern is possibly more acute with some of the other national institutions variables, such as the anti-self-dealing index. In tests not reported, we estimated instrumental variables (two-stage least squares) regression models. The national institutions variables are regressed on the common law indicator variable along with the world domestic IPO rate, country  $q$ , market cap / GDP, market turnover, log (GDP / capita), and world financial globalization in a first-stage regression. In the second-stage regression, the predicted values of the institutions variables are used instead of the actual values.<sup>13</sup> The results lead to similar conclusions as those that we report. Third, institutions affect the level of financial and economic development in a country. Consequently, we might be understating the influence of institutions. We estimated our regressions without the financial and economic development variables and found that the key inferences are unchanged. Fourth, one might be concerned that the dependent variable for counts or proceeds is censored at zero. We re-estimated our regressions using a Tobit model. The results in these regressions are often statistically more reliable than those reported in the table. In particular, the coefficients on rule of law and ownership concentration are now significant for the proceeds as well as for the count regressions. Fifth, we estimated regressions where the  $t$ -statistics are computed from standard errors that are double-clustered on both years and countries. The results are similar to those that we report. We also found consistent results for regressions that eliminate privatizations and count Hong Kong IPOs of Chinese companies as domestic IPOs.

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<sup>13</sup> DLLS (2008) use legal origins as an instrument for legal rules in a similar two-stage estimation procedure where the second stage explained financial development, such as IPO activity. LLS (2008, pp. 293-294) argue that such two-stage procedures are not recommended, however, since legal origins influence a broad range of rules and regulations and researchers cannot guarantee that the relevant ones are not omitted in the first stage. That is, legal rules can influence financial market outcomes other than through rules protecting investors.

In other unreported results, we investigated the “catching-up” hypothesis discussed earlier. We re-estimated the regressions of Table 4 adding a lagged dependent variable in the model. The catching-up hypothesis would imply a mean-reverting process in IPO activity so that a country with a high (low) IPO rate in the past would be expected to have a lower (higher) future IPO rate. We find no support for this hypothesis; if anything, we uncover evidence of positive autocorrelation in the series. More importantly, the statistical and economic significance of the institutional variables is unaffected when we add this lagged dependent variable.

Finally, we use our model to assess whether the U.S., in particular, has experienced fewer domestic IPOs than what would be predicted by the model, especially during the 2000s. To carry out this exercise, we re-estimate the regressions in both Panels of Table 4 with two added indicator variables. The first indicator variable takes a value of one for the U.S. in the 1990s and zero, otherwise (“U.S. dummy 1990s”); the second, takes a value of one for the U.S. in the 2000s, and zero, otherwise (“U.S. dummy 2000s”). In the base specification of Panel c of Table 4, we see that the U.S. has a 1.5% higher IPO rate than predicted by the model in the 1990s and a 1.94% lower rate in the 2000s. However, when we control for institutions, the positive coefficient for the indicator variable for the 1990s becomes mostly insignificant. In other words, given the quality of its institutions, the U.S. did not have more IPOs than predicted during the 1990s. Taking institutions into account, however, has the effect of making the IPO “deficit” in the 2000s even larger, at least for all but three of the institutions variables we study. For instance, when we include the common law dummy, the U.S. IPO deficit in the 2000s increases to 3.03%. This higher deficit translates into an additional 66 fewer IPOs per year in the U.S. during the 2000s.<sup>14</sup> When we turn to our regressions for domestic IPO proceeds in Panel d, the only reliable coefficient is a negative one associated with the U.S. dummy in the 2000s (-0.08%), implying again an IPO deficit of the U.S. in the 2000s. As for the regressions in Panel c, the IPO deficit in the U.S. becomes larger in

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<sup>14</sup> The average number of publicly-listed companies in any given year in the U.S. during the 2000s is 6002. The lower IPO rate of -1.94% from the base model specification implies 116 fewer IPOs per year in the U.S. during the 2000s, but the lower IPO rate of -3.03% from the specification with the common law dummy implies 182 fewer IPOs per year in the U.S. over that same time period. The implied difference is the 66 extra fewer IPOs per year.

magnitude once we account for the national institutions variables, with one exception. The higher deficit estimated in these specifications translates into as much as an additional \$14.8 billion per year less capital raised in domestic IPOs in the U.S. (using the specification with the investor protection variable).

We saw in Section II that the landscape of IPO activity changes dramatically during our sample period. In particular, the relative importance of U.S. IPOs and non-U.S. IPOs switched. The changes we discussed in Section II raise the question of whether the relation between IPO activity and institutions is stable through time and holds up with the rapid globalization of financial markets and with the rise of IPO activity all around the world. As explained in the introduction, there are good reasons to believe that globalization decreases the importance of national institutions.

### *C. Financial globalization and domestic IPO activity.*

Table 5 presents panel regression results of domestic IPO activity that are similar to those of Table 4, except that we interact our measure of world financial globalization from Lane and Milesi-Ferretti (2007) with the national institutions variables (“Institutions  $\times$  World financial globalization”). To preserve the interpretability of the coefficients (and to facilitate comparisons with Table 4), we de-mean the interacted variables across countries (each of the institutions variables, in turn) and across years (world financial globalization).<sup>15</sup> Indeed, the coefficients on the variables that comprise the base specification, including the world financial globalization measure, are very similar in magnitude and statistical significance to those we saw in Table 4.

In Panel a, we present the results for IPO counts as a fraction of the domestic number of listings and, in Panel b, for IPO proceeds as a fraction of GDP. In the first specification, which uses the common law dummy as the institutions variable, the coefficient on the interaction is reliably negative for domestic IPO counts (Panel a) and proceeds (Panel b). This weakening of the role of institutions as financial

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<sup>15</sup> Without de-meaning, the coefficient on the institutions variable in the regression with the interaction corresponds to the partial derivative of the dependent variable with respect to the institutions variable when the world financial globalization measure equals zero. This differs from the interpretation of the institutions variable in the regression without the interaction, where the interpretation of the coefficient on the institutions variable is the partial derivative when the world financial globalization measure is at its mean. By de-meaning, the interpretation of the coefficient on the institutions variable in the regression with the interaction is the same as that in regression without the interaction. See Ozer-Balli and Sorensen (2010).

globalization increases arises for just about every national institutions proxy variable that has a reliably positive coefficient in Table 4. In particular, financial globalization sharply reduces the positive impact of the anti-self-dealing, disclosure, burden of proof, public enforcement, and investor protection indexes on the domestic IPO rate. Consider the economic importance of this phenomenon for the example of the specification with the anti-self-dealing index. In Panel a of Table 4, we showed that a one-standard-deviation higher index score in anti-self-dealing (such as, from Switzerland to Canada) is associated with a 1.34% higher rate of domestic IPO activity. In Panel a of Table 5, we can compute how *different* the impact of such a higher score is in the 1990s when the level of financial globalization was lower (world financial globalization averaged 139% of world GDP) and in the 2000s when it was higher (it averaged 252% of world GDP). A one-standard-deviation increase is associated with a 1.56% higher rate of domestic IPO activity during the 1990s and only a 0.99% higher rate during the 2000s. At the peak of financial globalization in our sample period, none of the institutions variables have any impact on the domestic IPO rate measured using counts, except for a weak positive impact for the political risk variable. When we consider the IPO rate measured using proceeds in Panel b, none of the institutions variables have a material impact on the domestic IPO rate when financial globalization is at its peak.

It is interesting to note that the countries that were most important for IPO activity in the 1990s when the level of financial globalization was lower lost importance in the 2000s when the level was higher. The countries that were predominant in the 1990s, such as the U.S., U.K., Singapore, Australia, and Hong Kong, ceded substantial IPO share to a number of new countries with sizeable IPO activity in the 2000s, but with typically poorer institutions. This shift in IPO activity across countries is consistent with a decrease in the importance of institutions since, had institutions remained as important as back in the 1990s, the regressions would predict that the countries that gained prominence in IPO activity in the 2000s would not have done so.

In unreported regressions, we investigate the possibility that domestic sources of funding that substitute for IPOs became more available over time, which would explain the decrease in the role of institutions that we document. Instead of going public, pre-IPO private firms could be acquired so that

they would be able to invest more or so that insiders could cash out. Alternatively, firms could use debt finance instead of equity finance to fund their operational needs. To investigate these possibilities, we estimate regressions that add controls for the number of acquisitions of private firms in the country in the preceding year (scaled by the number of listed domestic firms, obtained from SDC) as well as for the level of private credit by deposit money banks and other financial institutions (as a fraction of GDP in the preceding year, obtained from the Financial Development and Structure database, originally used in Beck, Demirgüç-Kunt, and Levine (2000)). Including these control variables does not change our conclusions.

Our findings to now broadly support the law and finance theories we outlined in the introduction which predict a weakening role for national institutions for domestic IPO activity as financial globalization increases. Recall from SW (2002) that firms are larger, more valuable, and greater in number with higher dividends and less diversion of profits if legal protections are better. An important corollary of their model, however, is that more open capital markets are associated with greater IPO activity in a given country and that any differences in investor protection laws across countries will diminish in importance. More open markets enable firms to take advantage of financial development, the economic development, and the institutions of foreign countries. In the next section we provide further evidence in support of these theories. We document a negative relation between the rates of global IPO activity around the world and the national institutions variables. That is, countries with weaker institutions have higher rates of global IPO activity.

#### **IV. The consequences of financial globalization for global IPO activity around the world.**

Our evidence in Section II shows that IPOs in the rest of the world have become much more important and IPOs in the U.S. have become much less important. We have further revealed in Section III that a country's national institutions, whether corporate laws, securities laws, disclosure rules, or their enforcement in general, have become a less important factor for the extent of domestic IPO activity. One possible explanation is that countries' national institutions have become less important because firms wanting to pursue IPOs have found ever increasingly more ways made available by financial

globalization to avoid being hindered by institutional obstacles. These firms have gained greater and more advantageous access to global markets for their shares irrespective of the institutions of their country of domicile. For instance, firms can use global markets to go public to avoid being constrained by their home country.<sup>16</sup> Indeed, firms in many countries are pursuing IPOs in foreign markets or, at least, they are including international tranches in global IPOs. Is this pursuit related to the quality of national institutions? Is there evidence that this relationship is changing over time with greater financial globalization of markets?

*A. Understanding global IPO activity.*

In Table 6, we report the results of panel regressions that are similar to those of Table 4, except that our variable of interest is not domestic IPO counts and proceeds, but global IPO counts and proceeds. Global counts include foreign IPOs as well as global IPOs with a domestic and international tranche. Unlike our earlier analysis, these counts are deflated by the total number of IPOs, including both domestic and global IPOs, so we are evaluating how intensively the firms in a country pursue global opportunities. Global IPO proceeds include the U.S. dollar proceeds of the foreign IPOs as well as those from the international tranches of global IPOs (the proceeds assigned to the domestic tranche are excluded). Global IPO proceeds are deflated by total IPO proceeds, including domestic and global, so again we study the intensity of capital raising activity outside the home country. If firms from a given country do not have any IPOs, domestic or global, in a given year, the global IPO variables are set to missing.

The control variables are the same as those in Table 4 except that we refine the world domestic IPO rate for proceeds to include all domestic IPO proceeds, including proceeds raised in the domestic tranche of global IPOs. We also add the lagged level of domestic IPO activity in the country, the dependent variable in Table 4, to capture the extent to which the pursuit of global IPOs is a substitute or a complement activity (“Domestic IPO rate”). Slower domestic IPO activity may be associated with an increase in the intensity of global IPO activity because of the constraints imposed on emerging firms by

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<sup>16</sup> For 50 random global IPOs, we investigated whether firms that use global markets for their IPOs also tend to use foreign auditors, law firms, and underwriters. We found that this is the case, indicating that firms that have global IPOs also make use of monitors from outside of their country.

the existence of weak national institutions at home or simply better capital-raising opportunities elsewhere. But relatively underdeveloped capital markets at home may just as easily impose capacity constraints on the demands of capital-hungry firms such that high domestic IPO and global IPO activity arise together.

In the regressions exhibited in the first columns of Panel a for global IPO counts and of Panel b for global IPO proceeds, we find that the coefficient on the world domestic IPO rate is reliably positive and economically large. In the base model specification of Panel a, the coefficient of 1.95 implies that a one standard deviation increase in world domestic IPO activity worldwide is associated with a 4.49% increase in global IPO counts in a country, which represents 11.5% of the standard deviation of world domestic IPO activity. The equivalent coefficient for global IPO proceeds in Panel b is also significant and economically large. We also find reliable evidence that the level of domestic IPO activity is negatively related to the fraction of IPO counts and proceeds that are global. The economic importance of this relationship is even larger. For counts in Panel a, the coefficient on the domestic IPO rate is -3.34 which implies that a one standard deviation increase in domestic IPO counts per listed companies is associated with a 15.8% decrease in the fraction of IPOs that are global, which is about 41% of the standard deviation of global IPO activity. We also find that market turnover is negatively related to the intensity of global IPO activity by counts and proceeds. Both of these results are reliable indicators that robust domestic IPO activity is associated with fewer and less global IPO activity, not more. The coefficient on the world financial globalization measure is positive but not significant in the count regression and is positive and significant in the proceeds regression. None of the other variables add explanatory power, though the positive coefficient on country  $q$  is marginally significant for global IPO counts. The overall explanatory power of the base specification is reasonably good for the global IPO proceeds (adjusted  $R^2$  of 10%), and even better for global IPO counts (adjusted  $R^2$  of almost 22%).

*B. The importance of national institutions for global IPO activity.*

To the first regressions of both panels, we add one national institutions proxy variable in each subsequent column in both panels. We want to determine whether legal protections for minority investors,

securities laws, disclosure rules, and their enforcement in a country influence the intensity with which firms pursue global IPOs, even after controlling for the overall level of domestic IPO activity, growth opportunities, and market conditions. We find a reliable and important negative relationship for many of these variables. For example, countries with better anti-self-dealing measures are associated with much less global IPO activity. The negative coefficient on the anti-self-dealing index variable of -28.99 in Panel a implies a 15.05% lower fraction of global IPO counts, which accounts for about 39% of its standard deviation. We obtain a similarly reliable negative relationship for the intensity of global IPO counts using the common law dummy as well as the anti-director rights, disclosure, investor protection, and rule of law (at the 10% level) indexes. The political risk variable, the public enforcement and burden of proof indexes, are not significant. We expect a positive relationship between ownership and the extent of global IPO activity, and we confirm this in the last column of Panel a. The results for global IPO proceeds in Panel b are similar to those in Panel a. Again, while the institutions variables generally have positive and significant coefficients for domestic IPO proceeds, they have negative and significant coefficients for the global IPO proceeds regressions. As we saw in Panel a for the global IPO count results, the common law dummy, the anti-director rights (at the 10% level), anti-self-dealing, disclosure, investor protection, and rule of law indexes have reliably negative coefficients. The coefficient of -18.79 on the anti-self-dealing index implies a one standard deviation higher score (say, from that of Switzerland to Canada) is associated with a 4.58% decline in the fraction of IPO proceeds that are global offerings, which represents about 12% of its standard deviation. The ownership variable is positively related to the global fraction of IPO proceeds, as expected.

Though most of the institution proxies that have a positive and significant impact on domestic IPO activity in Table 4 have a negative and significant impact on global IPO activity in Table 6, there are some exceptions. For instance, the anti-director rights index is reliably significant in Table 6 but not in Table 4, while the public enforcement and burden of proof indexes have a strong positive coefficient in Table 4, but do not have a significant impact on global IPOs.

We saw in the previous section that national institutions became less important determinants of domestic IPO activity as globalization increased. We now explore whether it holds for global IPOs.

*C. Global IPO activity and financial globalization.*

If institutions have become less important for the level of domestic IPO activity and if financial globalization and the accessibility of global IPOs are related to this development, then we would expect that national institutions have become similarly less important for the intensity of global IPO activity as financial globalization increased. Panels a and b in Table 7 represent the equivalent tests to those in Table 5, but for the fraction of total IPOs that arise in global form by count and by proceeds, respectively. We employ the same base specification for our panel regressions as in Table 6, but we introduce an interaction variable for the institutions variables with the world financial globalization measure. As in Table 5, we de-mean the variables that constitute the interaction variable (across countries for the institutions proxies and across years for world financial globalization) in order to preserve the interpretability of the coefficients on the variables in the base specification.

In Panel a, we find that some institution variables do become less important as financial globalization increases, but not all of them. With greater financial globalization, the common law dummy, the anti-director rights and the anti-self-dealing indexes lose their importance. However, the importance of the disclosure, investor protection, and rule of law indexes are not affected by financial globalization. Finally, ownership concentration is also not affected by financial globalization. We can capture how much the economic magnitude of the relationship between the institutions variables and global IPO activity rates changes with increased financial globalization using the estimates in the respective models for the anti-self-dealing variable (e.g., Model (4) in the count regressions of Panel a in both Table 6 and 7). On average, a one-standard-deviation higher score on the anti-self-dealing index is associated with a 6.67% lower fraction of global IPOs (on average, the global fraction by counts is just over 44%). During the 1990s, when the level of financial globalization is lower (world financial globalization averages 139% of world GDP), the same one-standard-deviation higher score on anti-self-dealing is associated with a 8.27% lower fraction of global IPOs, whereas this difference is associated with only a 5.03% lower fraction

during the 2000s (when world financial globalization averages 252% of world GDP). It seems that the consequence of greater financial globalization for global IPO activity is somewhat less dramatic than for domestic IPO activity. The results for proceeds in Panel b are similar except that the interaction is not significant for the anti-director index.

What is the bottom-line of our findings on global IPOs? As we would expect, global IPOs are a way for firms to exploit the better institutions of foreign countries to have a successful or more profitable IPO. The advantage of the institutions of foreign countries is inversely related to the quality of a firm's domestic institutions, so that it is not surprising that domestic institutions play an opposite role for global and domestic IPOs. However, while we find evidence for both domestic and global IPOs that domestic institutions become less important in the 2000s than in the 1990s, this evidence is substantially stronger for domestic IPOs than it is for global IPOs. A plausible explanation for this finding is that financial globalization increasingly enables firms whose value is most closely tied to the quality of institutions to use global IPOs and to take advantage of the institutions of foreign countries. As a result, firms that use domestic IPOs are firms for which the quality of institutions is relatively less important.

## **V. Conclusions.**

This paper documents dramatic changes in the IPO landscape around the world. U.S. IPOs and IPOs from other common law countries have become less important, whether one looks at counts or at proceeds. In fact, U.S. IPO activity has generally not kept pace with the economic importance of the U.S.

Financial globalization plays a critical role in the increasing importance of IPOs by non-U.S. firms. Financial globalization makes it easier for firms to have global IPOs which enable them to mitigate the impact of weaker institutions in their home country. Greater globalization also affects the monitoring of firms as holdings by institutional investors grow throughout the world. With greater financial globalization, a country's institutions become less important for the domestic IPO rate. In addition, greater financial globalization is associated with an increase in the rate of global IPOs. Global IPOs have played a critical role in increasing the importance of IPOs by non-U.S. firms. Though firms in countries

with weaker institutions are less likely to go public with a domestic IPO, they are more likely to go public with a global IPO. That is, global IPOs enable firms to overcome poor institutions in their country of origin. Perhaps as a result, the laws and institutions of a firm's country of origin have become significantly less important in affecting the rate and pace of IPO activity in a country.

There are important global drivers in domestic IPO activity. Higher levels of worldwide IPO activity outside a country are strongly and positively related to the level of IPO activity in that country. However, IPO activity is also related to domestic market conditions. Firms are more likely to choose to go public at home when valuations are higher in the home market.

Our paper leaves open some important issues. First, although we find clear evidence that institutions have become less important in affecting a country's IPO activity, it could be that laws and regulations that we do not account for affect IPO activity. Further work should therefore examine the impact of changes in laws that are not captured by our institutional proxy variables. Second, we do not investigate the impact of financial globalization on individual IPOs. An investigation of the extent to which firms going public in financially open countries make use of institutions and resources from other countries would help in understanding better the impact of financial globalization on IPO activity. Finally, our focus is resolutely on cross-country variation in IPO activity, but as a result we highlight the decreasing role of IPOs in the U.S. in the 2000s. Further work should address that decrease and explain it. Much recent research and policy debates have focused on competition between London and New York. We showed in earlier research (Doidge, Karolyi, and Stulz, 2009) that New York was not losing market share to London in attracting secondary listings of foreign firms. The global financial crisis in 2008-2009 has made this issue largely obsolete. However, this paper shows that focusing on the regulatory advantages of London versus New York misses the big picture. To abuse once more Thomas Friedman's wonderful analogy, the IPO world is clearly becoming flat.<sup>17</sup>

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<sup>17</sup> See Thomas L. Friedman, *The World is Flat: A Brief History of the Twenty-First Century*.

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**Table 1. The IPO sample: 1990 to 2007.**

The initial sample includes 38,724 observations from 1990 to 2007 that SDC identifies as an IPO. IPOs with a single domestic tranche flagged as a private placement, global offers with tranches that have issue dates 30 or more days apart, transactions that do not contain any information on proceeds raised or SIC codes, and IPOs by REITs and investment funds are excluded. IPOs where the country of origin has no data and IPOs from countries where there were no domestic IPOs (only global IPOs) during the sample period are also excluded. SDC records data for some IPOs over multiple lines. These observations are consolidated into one line. The final sample includes 29,361 IPOs from 89 countries. Of these, 24,122 are domestic IPOs and 5,239 are global IPOs (IPOs in which some or all of the shares are sold outside the home country of the firm going public). Panel a shows IPO counts and Panel b shows IPOs proceeds. Domestic IPO proceeds do not include proceeds raised in the domestic tranche of global IPOs. For global IPOs, the panel reports total proceeds raised in global IPOs (proceeds raised in the domestic and international tranches) and global proceeds raised in global IPOs (proceeds raised in the international tranches only). Proceeds are in constant 2007 U.S. dollars (billions).

Panel a. IPO counts.			
Year	All IPOs	Domestic IPOs	Global IPOs
1990	303	248	55
1991	891	804	87
1992	1,339	1,211	128
1993	2,078	1,860	218
1994	2,739	2,474	265
1995	2,688	2,433	255
1996	3,100	2,766	334
1997	1,959	1,580	379
1998	1,232	922	310
1999	1,589	1,006	583
2000	2,117	1,452	665
2001	971	798	173
2002	914	809	105
2003	910	809	101
2004	1,529	1,297	232
2005	1,473	1,223	250
2006	1,679	1,314	365
2007	1,850	1,116	734
Total	29,361	24,122	5,239

**Table 1, continued.**

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Panel b. IPO proceeds.

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Year	All IPOs	Domestic IPOs	Global IPOs: total	Global IPOs: global only
1990	\$29.6	\$18.5	\$11.1	\$8.8
1991	\$71.7	\$37.9	\$33.8	\$20.5
1992	\$60.8	\$35.5	\$25.3	\$10.6
1993	\$150.2	\$92.1	\$58.2	\$28.8
1994	\$157.7	\$77.7	\$80.0	\$43.2
1995	\$116.4	\$47.1	\$69.3	\$37.3
1996	\$168.8	\$81.7	\$87.1	\$45.2
1997	\$179.8	\$69.8	\$110.0	\$49.2
1998	\$138.2	\$32.6	\$105.6	\$39.8
1999	\$210.0	\$59.3	\$150.7	\$63.2
2000	\$242.2	\$51.8	\$190.4	\$94.0
2001	\$108.1	\$35.7	\$72.4	\$32.1
2002	\$76.5	\$46.7	\$29.7	\$13.4
2003	\$59.1	\$34.8	\$24.3	\$15.2
2004	\$133.8	\$62.2	\$71.6	\$45.1
2005	\$149.4	\$82.6	\$66.8	\$52.4
2006	\$223.7	\$121.6	\$102.1	\$89.8
2007	\$278.6	\$89.9	\$188.7	\$169.4
Total	\$2,554.6	\$1,077.5	\$1,477.1	\$858.1

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**Table 2. IPO activity for the top 25 countries around the world: 1990 to 2007.**

IPO data is from SDC and includes 29,361 IPOs from 89 countries over the period from 1990 to 2007. Panel a lists the top 25 countries based on total IPO counts. Panel b lists the top 25 countries based on total IPO proceeds. Domestic IPO proceeds do not include proceeds raised in the domestic tranche of global IPOs. For global IPOs the panel reports total proceeds raised in global IPOs (proceeds raised in the domestic and international tranches) and global proceeds raised in global IPOs (proceeds raised in the international tranches only). Proceeds are in constant 2007 U.S. dollars (billions).

Panel a. IPO counts.			
Country	All IPOs	Domestic IPOs	Global IPOs
United States	6,126	4,931	1,195
India	4,867	4,777	90
Japan	2,234	2,130	104
Canada	2,225	2,020	205
China	1,764	1,300	464
United Kingdom	1,650	1,356	294
Australia	1,558	1,400	158
Hong Kong	822	541	281
Taiwan	822	808	14
South Korea	779	752	27
France	750	503	247
Malaysia	722	697	25
Germany	573	288	285
Singapore	488	404	84
Thailand	408	333	75
Indonesia	273	189	84
Pakistan	249	247	2
Italy	244	54	190
Greece	185	148	37
Norway	179	123	56
Poland	175	133	42
Israel	155	13	142
Sweden	143	53	90
Brazil	128	60	68
Netherlands	120	26	94
<i>Total: top 25</i>	<i>27,639</i>	<i>23,286</i>	<i>4,353</i>
<i>Rest of world</i>	<i>1722</i>	<i>836</i>	<i>886</i>
<i>Total: all countries</i>	<i>29,361</i>	<i>24,122</i>	<i>5,239</i>

**Table 2, continued.**

Country	Panel b. IPO proceeds.			
	All IPOs	Domestic IPOs	Global IPOs: total	Global IPOs: global only
United States	\$647.7	\$352.3	\$295.4	\$61.6
China	\$254.6	\$110.1	\$144.5	\$133.1
Japan	\$204.1	\$135.2	\$68.9	\$22.1
United Kingdom	\$196.3	\$77.1	\$119.2	\$68.9
France	\$122.3	\$9.7	\$112.6	\$54.3
Germany	\$106.6	\$27.6	\$79.0	\$45.0
Italy	\$84.2	\$9.7	\$74.5	\$32.4
Australia	\$76.3	\$34.4	\$41.9	\$18.8
Canada	\$68.6	\$47.7	\$20.9	\$15.2
Hong Kong	\$63.6	\$12.9	\$50.7	\$43.6
South Korea	\$58.2	\$46.1	\$12.1	\$10.5
Russian Federation	\$43.6	\$13.9	\$29.7	\$29.7
Spain	\$41.5	\$3.2	\$38.3	\$18.4
Netherlands	\$39.6	\$4.1	\$35.5	\$28.2
Brazil	\$39.3	\$14.9	\$24.4	\$23.3
Switzerland	\$37.1	\$9.6	\$27.5	\$20.2
Sweden	\$33.9	\$3.4	\$30.5	\$17.3
India	\$32.2	\$17.8	\$14.4	\$12.5
Taiwan	\$27.1	\$25.5	\$1.6	\$1.5
Bermuda	\$26.5	\$0.1	\$26.4	\$26.4
Thailand	\$22.9	\$11.0	\$11.9	\$6.5
Singapore	\$20.3	\$7.9	\$12.4	\$10.5
Indonesia	\$20.3	\$5.0	\$15.3	\$9.9
Mexico	\$19.6	\$7.0	\$12.5	\$10.2
Norway	\$18.6	\$6.7	\$11.9	\$8.6
<i>Total: top 25</i>	<i>\$2,305.1</i>	<i>\$992.8</i>	<i>\$1312.3</i>	<i>\$728.7</i>
<i>Rest of world</i>	<i>\$249.5</i>	<i>\$84.7</i>	<i>\$164.8</i>	<i>\$129.4</i>
<i>Total: all countries</i>	<i>\$2,554.6</i>	<i>\$1,077.5</i>	<i>\$1,477.1</i>	<i>\$858.1</i>

**Table 3. Domestic IPO activity: 1990 to 2007.**

IPO data is from SDC and includes 23,907 domestic IPOs from 54 countries that have data available for GDP and country  $q$  for at least one year during the sample period from 1990 to 2007. For each country, domestic IPO counts and proceeds are summed annually. Panel a shows annual summary statistics for domestic IPO counts scaled by the lagged number of domestic firms. Panel b shows annual summary statistics for domestic IPO proceeds scaled by lagged GDP. Both measures are multiplied by 100. Domestic IPO proceeds do not include proceeds from the domestic tranche of global IPOs. Both measures of domestic IPO activity are winsorized at the 1<sup>st</sup> and 99<sup>th</sup> percentiles. Country-years with no data for the number of domestic firms or GDP are excluded.

	1990	1991	1992	1993	1994	1995	1996	1997	1998	1999	2000	2001	2002	2003	2004	2005	2006	2007
Panel a. Domestic IPO counts scaled by the lagged number of domestic firms.																		
Mean	0.429	2.835	4.017	4.463	5.385	2.798	3.346	3.262	1.868	2.324	2.894	1.625	1.559	1.477	2.092	2.125	2.263	2.183
Median	0.000	0.000	0.784	0.905	1.858	1.075	0.976	0.770	0.566	0.363	0.847	0.000	0.295	0.000	0.405	1.090	1.376	1.207
Std deviation	0.938	5.250	6.244	6.583	6.871	4.658	5.406	5.274	3.017	3.472	4.197	2.583	2.424	2.667	3.285	2.897	2.448	3.315
Minimum	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Maximum	3.704	25.234	25.234	25.234	25.234	25.234	25.234	25.234	12.827	15.164	16.620	10.923	10.103	14.734	13.752	13.855	10.043	19.476
# of countries with zero IPOs	41	33	25	21	19	23	23	17	21	24	19	29	25	28	19	16	13	13
Panel b. Domestic IPO proceeds scaled by lagged GDP.																		
Mean	0.077	0.148	0.161	0.226	0.261	0.126	0.145	0.156	0.072	0.171	0.121	0.066	0.072	0.075	0.084	0.179	0.195	0.192
Median	0.000	0.000	0.005	0.058	0.018	0.011	0.009	0.017	0.007	0.016	0.024	0.000	0.002	0.000	0.003	0.111	0.135	0.059
Std deviation	0.272	0.277	0.309	0.390	0.379	0.217	0.257	0.301	0.136	0.387	0.239	0.196	0.136	0.159	0.137	0.214	0.258	0.365
Minimum	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Maximum	1.521	1.186	1.486	1.521	1.363	1.050	1.149	1.239	0.626	1.521	1.330	1.332	0.670	0.640	0.517	0.834	1.521	1.521
# of countries with zero IPOs	39	30	25	20	19	23	23	17	21	24	19	29	25	28	19	16	13	13

**Table 4. Determinants of domestic IPO activity: 1990 to 2007.**

The dependent variable is each country's annual measure of domestic IPO activity. IPO data is from SDC and includes 23,907 domestic IPOs from 54 countries that have data available for GDP and country  $q$  for at least one year during the sample period from 1990 to 2007. For each country, domestic IPO counts and proceeds are summed annually. Panel a shows regressions where the dependent variable is each country's annual domestic IPO count scaled by the lagged number of domestic firms. Panel b shows regressions where the dependent variable is each country's annual domestic IPO proceeds scaled by lagged GDP. Domestic IPO proceeds do not include proceeds from the domestic tranche of global IPOs. Both measures of domestic IPO activity are multiplied by 100 and are winsorized at the 1<sup>st</sup> and 99<sup>th</sup> percentiles. Country-years with no data for the number of domestic firms or GDP are excluded. In Panels a and c (Panels b and d), the world domestic IPO rate is based on counts (proceeds). With the exception of the institutions variables, all variables are lagged by one year. Variables are defined in Appendix B. The  $t$ -statistics (in parentheses) are adjusted for clustering on countries – they are computed assuming observations are independent across countries, but not within countries. \*, \*\*, and \*\*\* indicate statistical significance at the 10%, 5%, and 1% levels, respectively.

Panel a. Domestic IPO counts scaled by the lagged number of domestic firms.											
		Common law	Anti-director	Anti-self-dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	0.430 (0.17)	-1.076 (-0.38)	-0.573 (-0.14)	-3.225 (-1.49)	-6.831*** (-2.89)	-2.699 (-1.18)	-5.256** (-2.15)	-4.598* (-1.89)	-0.040 (-0.02)	5.195 (1.42)	4.187 (1.55)
Institutions variable		1.548* (1.83)	0.207 (0.46)	5.602*** (4.22)	7.177*** (4.84)	2.668* (1.94)	3.821** (2.32)	4.099*** (3.00)	0.063 (1.32)	1.145** (2.16)	-7.222*** (-2.80)
World domestic IPO rate	0.302*** (3.77)	0.296*** (3.57)	0.298** (3.53)	0.292** (3.54)	0.229** (3.28)	0.227*** (3.40)	0.223** (3.24)	0.224*** (3.22)	0.307*** (3.75)	0.323*** (3.89)	0.229*** (3.45)
Country $q$	2.703*** (3.13)	2.732*** (3.30)	2.816*** (3.29)	2.792*** (3.18)	2.633** (2.68)	2.644*** (2.72)	2.735** (2.59)	2.657*** (2.70)	2.579*** (3.10)	2.336*** (2.86)	2.581** (2.64)
Market cap / GDP	1.068* (1.97)	0.542 (1.03)	0.926 (1.64)	0.003 (0.01)	-0.157 (-0.31)	0.850 (1.48)	0.769 (1.53)	0.499 (0.91)	0.995* (1.81)	0.935* (1.75)	1.052* (1.72)
Market turnover	1.158** (2.14)	1.166* (1.92)	1.132* (1.93)	1.107** (2.11)	0.497 (0.86)	0.700 (1.30)	0.818 (1.48)	0.730 (1.26)	1.169** (2.20)	1.126* (1.95)	0.271 (0.52)
Log (GDP / capita)	-0.273 (-0.92)	-0.140 (-0.42)	-0.250 (-0.76)	-0.154 (-0.62)	0.136 (0.64)	-0.005 (-0.02)	0.203 (0.84)	0.141 (0.58)	-0.732 (-1.34)	-0.936* (-2.00)	-0.254 (-1.01)
World financial globalization	-0.655** (-2.02)	-0.571* (-1.81)	-0.625* (-1.93)	-0.413 (-1.43)	-0.356 (-1.23)	-0.602* (-1.96)	-0.665** (-2.37)	-0.548* (-1.91)	-0.613* (-1.79)	-0.262 (-0.71)	-0.501 (-1.42)
Number of observations	890	890	890	890	777	777	777	777	890	890	777
Adjusted R <sup>2</sup>	0.1358	0.1568	0.1371	0.2107	0.2295	0.1561	0.1702	0.1769	0.1452	0.1493	0.1695

**Table 4, continued.**

Panel b. Domestic IPO proceeds scaled by lagged GDP.											
		Common law	Anti-director	Anti-self-dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	0.046 (0.34)	-0.061 (-0.42)	-0.074 (-0.38)	-0.178 (-1.47)	-0.272** (-2.10)	-0.089 (-0.67)	-0.242 (-1.40)	-0.181 (-1.21)	0.039 (0.30)	0.199 (0.94)	0.093 (0.60)
Institutions variable		0.108** (2.63)	0.024 (1.11)	0.340*** (4.65)	0.315*** (3.25)	0.110* (1.81)	0.204** (2.13)	0.186** (2.64)	0.001 (0.39)	0.036 (1.21)	-0.165 (-1.26)
World domestic IPO rate	0.634*** (4.90)	0.624*** (4.67)	0.613*** (4.60)	0.622*** (4.69)	0.576*** (4.20)	0.593*** (4.41)	0.580*** (4.32)	0.587*** (4.34)	0.640*** (4.87)	0.653*** (4.99)	0.592*** (4.38)
Country $q$	0.158*** (3.10)	0.160*** (3.05)	0.171*** (3.28)	0.163*** (2.96)	0.167** (2.47)	0.167** (2.66)	0.172** (2.51)	0.167** (2.54)	0.156*** (3.05)	0.147*** (2.71)	0.165** (2.58)
Market cap / GDP	0.105*** (3.12)	0.069** (2.32)	0.089** (2.80)	0.041 (1.43)	0.049 (1.55)	0.095*** (2.72)	0.086** (2.59)	0.077** (2.21)	0.104*** (3.13)	0.101*** (3.08)	0.107*** (2.85)
Market turnover	0.059* (1.83)	0.060 (1.67)	0.056* (1.71)	0.056* (1.80)	0.031 (0.98)	0.041 (1.24)	0.046 (1.26)	0.042 (1.15)	0.059* (1.84)	0.058* (1.78)	0.032 (1.01)
Log (GDP / capita)	-0.022 (-1.60)	-0.013 (-0.86)	-0.020 (-1.24)	-0.015 (-1.26)	-0.006 (-0.47)	-0.012 (-0.99)	-0.001 (-0.05)	-0.005 (-0.41)	-0.029 (-1.18)	-0.043* (-1.80)	-0.018 (-1.57)
World financial globalization	-0.053** (-2.26)	-0.046** (-2.01)	-0.048** (-2.12)	-0.037* (-1.79)	-0.041* (-1.72)	-0.052** (-2.07)	-0.053** (-2.19)	-0.049* (-1.99)	-0.052** (-2.23)	-0.042* (-1.75)	-0.052* (-1.89)
Number of observations	890	890	890	890	777	777	777	777	890	890	777
Adjusted R <sup>2</sup>	0.1281	0.1538	0.1353	0.1967	0.1652	0.1343	0.1485	0.1452	0.1277	0.1307	0.1297

**Table 4, continued.**

Panel c. Domestic IPO counts scaled by the lagged number of domestic firms.											
		Common law	Anti-director	Anti-self-dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	0.341 (0.13)	-1.404 (-0.47)	-0.631 (-0.16)	-3.399 (-1.53)	-7.748*** (-3.09)	-3.162 (-1.29)	-6.131** (-2.36)	-5.608** (-2.20)	-0.091 (-0.04)	5.099 (1.38)	4.193 (1.53)
Institutions variable		1.631* (1.82)	0.205 (0.45)	5.613*** (4.16)	7.733*** (4.95)	2.928* (1.91)	4.272** (2.43)	4.748*** (3.29)	0.063 (1.31)	1.138** (2.14)	-7.653*** (-2.79)
World domestic IPO rate	0.303*** (3.78)	0.294*** (3.56)	0.298*** (3.54)	0.292*** (3.55)	0.226*** (3.27)	0.224*** (3.42)	0.220*** (3.23)	0.220*** (3.21)	0.307*** (3.76)	0.324*** (3.90)	0.227*** (3.46)
Country $q$	2.702*** (3.11)	2.776*** (3.35)	2.809*** (3.27)	2.815*** (3.18)	2.757*** (2.79)	2.717*** (2.78)	2.843** (2.64)	2.782*** (2.83)	2.570*** (3.08)	2.332*** (2.84)	2.651*** (2.70)
Market cap / GDP	1.085* (1.99)	0.545 (1.03)	0.942 (1.65)	0.024 (0.05)	-0.217 (-0.43)	0.847 (1.46)	0.755 (1.48)	0.431 (0.79)	1.009* (1.82)	0.950* (1.77)	1.075* (1.76)
Market turnover	1.187** (2.17)	1.215* (1.99)	1.159* (1.94)	1.145** (2.17)	0.541 (0.97)	0.743 (1.40)	0.883 (1.64)	0.790 (1.39)	1.192** (2.22)	1.151* (1.97)	0.289 (0.57)
Log (GDP / capita)	-0.275 (-0.92)	-0.119 (-0.35)	-0.254 (-0.77)	-0.148 (-0.59)	0.185 (0.86)	0.016 (0.06)	0.256 (1.05)	0.202 (0.82)	-0.735 (-1.33)	-0.936* (-1.99)	-0.248 (-0.98)
World financial globalization	-0.614* (-1.85)	-0.542* (-1.69)	-0.583* (-1.77)	-0.383 (-1.29)	-0.326 (-1.11)	-0.567* (-1.81)	-0.641** (-2.23)	-0.512* (-1.75)	-0.570 (-1.63)	-0.223 (-0.60)	-0.469 (-1.32)
U.S. dummy 1990s	1.500** (2.61)	0.297 (0.32)	1.609*** (2.74)	0.755 (1.25)	-1.405 (-1.63)	-0.094 (-0.08)	-0.566 (-0.55)	-1.191 (-1.10)	1.700*** (2.83)	1.586*** (2.74)	-0.233 (-0.26)
U.S. dummy 2000s	-1.942** (-2.53)	-3.027*** (-3.00)	-1.787** (-2.09)	-2.389*** (-3.41)	-4.193*** (-4.59)	-3.332*** (-2.91)	-3.903*** (-3.59)	-4.382*** (-3.98)	-1.637* (-1.99)	-1.729** (-2.11)	-3.190*** (-3.45)
Number of observations	890	890	890	890	777	777	777	777	890	890	777
Adjusted R <sup>2</sup>	0.1368	0.1590	0.1381	0.2118	0.2389	0.1604	0.1769	0.1863	0.1462	0.1502	0.1734

**Table 4, continued.**

Panel d. Domestic IPO proceeds scaled by lagged GDP.											
		Common law	Anti-director	Anti-self-dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	0.038 (0.28)	-0.090 (-0.62)	-0.077 (-0.40)	-0.194 (-1.58)	-0.329** (-2.51)	-0.120 (-0.89)	-0.306* (-1.74)	-0.246 (-1.68)	0.033 (0.25)	0.189 (0.89)	0.100 (0.64)
Institutions variable		0.118*** (2.74)	0.024 (1.09)	0.344*** (4.64)	0.354*** (3.46)	0.132* (1.95)	0.240** (2.43)	0.232*** (3.35)	0.001 (0.36)	0.035 (1.19)	-0.195 (-1.44)
World domestic IPO rate	0.629*** (4.85)	0.611*** (4.56)	0.610*** (4.57)	0.615*** (4.62)	0.557*** (4.08)	0.580*** (4.33)	0.561*** (4.19)	0.568*** (4.23)	0.635*** (4.82)	0.648*** (4.95)	0.582*** (4.28)
Country $q$	0.160*** (3.13)	0.165*** (3.15)	0.172*** (3.28)	0.167*** (3.01)	0.175** (2.59)	0.172*** (2.77)	0.180** (2.59)	0.176* (2.67)	0.158*** (3.07)	0.148*** (2.73)	0.170** (2.66)
Market cap / GDP	0.106*** (3.13)	0.067** (2.26)	0.089*** (2.81)	0.041 (1.43)	0.044 (1.38)	0.093** (2.68)	0.083** (2.52)	0.071** (2.06)	0.105*** (3.14)	0.102*** (3.10)	0.107*** (2.86)
Market turnover	0.060* (1.85)	0.062* (1.74)	0.057* (1.72)	0.058* (1.85)	0.033 (1.04)	0.042 (1.30)	0.049 (1.35)	0.044 (1.22)	0.061* (1.86)	0.059* (1.80)	0.032 (1.00)
Log (GDP / capita)	-0.022 (-1.55)	-0.011 (-0.69)	-0.019 (-1.22)	-0.014 (-1.15)	-0.002 (-0.19)	-0.010 (-0.82)	0.004 (0.30)	-0.001 (-0.07)	-0.028 (-1.12)	-0.042* (-1.76)	-0.017 (-1.52)
World financial globalization	-0.052** (-2.21)	-0.046* (-1.99)	-0.047** (-2.08)	-0.037* (-1.77)	-0.040* (-1.69)	-0.052** (-2.04)	-0.053** (-2.20)	-0.048* (-1.96)	-0.052** (-2.19)	-0.042* (-1.72)	-0.052* (-1.86)
U.S. dummy 1990s	-0.019 (-0.60)	-0.106** (-2.20)	-0.007 (-0.23)	-0.065* (-1.98)	-0.162*** (-2.92)	-0.101* (-1.78)	-0.143** (-2.35)	-0.160*** (-2.80)	-0.017 (-0.50)	-0.017 (-0.51)	-0.076* (-1.91)
U.S. dummy 2000s	-0.080** (-2.09)	-0.159*** (-3.17)	-0.062 (-1.65)	-0.107*** (-2.93)	-0.184*** (-3.70)	-0.143** (-2.53)	-0.190*** (-3.41)	-0.199*** (-3.98)	-0.076* (-1.93)	-0.073* (-1.89)	-0.115** (-2.52)
Number of observations	890	890	890	890	777	777	777	777	890	890	777
Adjusted R <sup>2</sup>	0.1269	0.1560	0.1337	0.1967	0.1708	0.1359	0.1533	0.1508	0.1264	0.1294	0.1299

**Table 5. Determinants of domestic IPO activity and financial globalization.**

The dependent variable is each country's annual measure of domestic IPO activity. IPO data is from SDC and includes 23,907 domestic IPOs from 54 countries that have data available for GDP and country  $q$  for at least one year during the sample period from 1990 to 2007. For each country, domestic IPO counts and proceeds are summed annually. Panel a shows regressions where the dependent variable is each country's annual domestic IPO count scaled by the lagged number of domestic firms. Panel b shows regressions where the dependent variable is each country's annual domestic IPO proceeds scaled by lagged GDP. Domestic IPO proceeds do not include proceeds from the domestic tranche of global IPOs. Both measures of domestic IPO activity are multiplied by 100 and are winsorized at the 1<sup>st</sup> and 99<sup>th</sup> percentiles. Country-years with no data for the number of domestic firms or GDP are excluded. In Panel a (Panel b), the world domestic IPO rate is based on counts (proceeds). With the exception of the institutions variables, all variables are lagged by one year. Variables are defined in Appendix B. The  $t$ -statistics (in parentheses) are adjusted for clustering on countries – they are computed assuming observations are independent across countries, but not within countries. \*, \*\*, and \*\*\* indicate statistical significance at the 10%, 5%, and 1% levels, respectively.

	Panel a. Domestic IPO counts scaled by the lagged number of domestic firms.									
	Common law	Anti-director	Anti-self-dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	-1.678 (-0.63)	-1.070 (-0.39)	-1.264 (-0.61)	-2.788 (-1.46)	-2.548 (-1.36)	-4.328** (-2.22)	-3.552* (-1.80)	3.780 (0.80)	5.544 (1.33)	-0.059 (-0.03)
Institutions variable	1.611* (1.93)	0.220 (0.47)	5.672*** (4.24)	7.048*** (4.74)	2.669* (1.93)	3.819** (2.32)	4.095*** (2.96)	0.063 (1.31)	1.115** (2.10)	-7.252*** (-2.82)
Institutions × World financial globalization	-1.958** (-2.60)	-0.231 (-0.58)	-2.515** (-2.36)	-4.254*** (-3.22)	-2.114* (-1.94)	-2.789** (-2.06)	-3.568** (-2.64)	0.049* (1.87)	0.225 (0.67)	-0.918 (-0.43)
World domestic IPO rate	0.299*** (3.55)	0.300*** (3.55)	0.294*** (3.53)	0.239*** (3.37)	0.228*** (3.39)	0.228*** (3.27)	0.230*** (3.27)	0.320*** (3.89)	0.324*** (3.90)	0.229*** (3.45)
Country $q$	2.647*** (3.32)	2.796*** (3.32)	2.714*** (3.13)	2.470** (2.59)	2.600*** (2.72)	2.568** (2.44)	2.495** (2.60)	2.417*** (3.05)	2.298*** (2.87)	2.577** (2.63)
Market cap / GDP	0.572 (1.09)	0.925 (1.63)	0.030 (0.07)	-0.046 (-0.09)	0.882 (1.53)	0.820 (1.61)	0.568 (1.03)	0.965* (1.76)	0.913* (1.69)	1.046* (1.71)
Market turnover	1.236** (2.13)	1.154** (2.02)	1.108** (2.11)	0.536 (0.96)	0.724 (1.37)	0.816 (1.48)	0.779 (1.40)	1.185** (2.32)	1.115* (1.94)	0.257 (0.49)
Log (GDP / capita)	-0.138 (-0.42)	-0.247 (-0.74)	-0.150 (-0.61)	0.119 (0.56)	-0.004 (-0.02)	0.197 (0.82)	0.140 (0.58)	-0.756 (-1.38)	-0.924* (-1.99)	-0.256 (-1.02)
World financial globalization	-0.586** (-2.03)	-0.618* (-1.95)	-0.397 (-1.41)	-0.363 (-1.27)	-0.602* (-1.99)	-0.659** (-2.52)	-0.545* (-1.98)	-0.614* (-1.90)	-0.266 (-0.72)	-0.501 (-1.44)
Number of observations	890	890	890	777	777	777	777	890	890	777
Adjusted R <sup>2</sup>	0.1727	0.1375	0.2175	0.2483	0.1622	0.1788	0.1932	0.1513	0.1493	0.1688

**Table 5, continued.**

	Panel b. Domestic IPO proceeds scaled by lagged GDP.									
	Common law	Anti-director	Anti-self-dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	-0.114 (-0.86)	-0.086 (-0.60)	-0.081 (-0.71)	-0.138 (-1.14)	-0.135 (-1.18)	-0.226* (-1.74)	-0.176 (-1.41)	0.005 (0.02)	0.151 (0.65)	-0.089 (-0.71)
Institutions variable	0.113*** (2.79)	0.027 (1.18)	0.347*** (4.81)	0.306*** (3.18)	0.110* (1.81)	0.204** (2.14)	0.186** (2.68)	0.001 (0.39)	0.040 (1.34)	-0.160 (-1.22)
Institutions × World financial globalization	-0.135*** (-2.87)	-0.043** (-2.04)	-0.259*** (-2.92)	-0.282** (-2.60)	-0.222*** (-2.96)	-0.152* (-1.69)	-0.265*** (-2.97)	-0.000 (-0.06)	-0.030 (-1.44)	0.144 (1.20)
World domestic IPO rate	0.631*** (4.74)	0.625*** (4.61)	0.633*** (4.73)	0.596*** (4.25)	0.600*** (4.42)	0.584*** (4.32)	0.600*** (4.37)	0.639*** (4.82)	0.650*** (4.92)	0.597*** (4.43)
Country $q$	0.154*** (3.02)	0.168*** (3.25)	0.155*** (2.89)	0.156** (2.35)	0.162** (2.65)	0.163** (2.41)	0.155** (2.41)	0.157*** (3.06)	0.152*** (2.83)	0.166** (2.57)
Market cap / GDP	0.071** (2.50)	0.088*** (2.89)	0.044 (1.67)	0.057* (1.76)	0.098*** (2.83)	0.088** (2.65)	0.082** (2.40)	0.104*** (3.12)	0.104*** (3.14)	0.108*** (2.89)
Market turnover	0.065* (1.92)	0.060* (1.86)	0.056* (1.80)	0.034 (1.07)	0.043 (1.31)	0.046 (1.27)	0.045 (1.30)	0.059* (1.84)	0.059* (1.77)	0.035 (1.08)
Log (GDP / capita)	-0.013 (-0.87)	-0.019 (-1.19)	-0.015 (-1.28)	-0.007 (-0.57)	-0.012 (-0.98)	-0.001 (-0.08)	-0.005 (-0.42)	-0.029 (-1.17)	-0.045* (-1.85)	-0.017 (-1.53)
World financial globalization	-0.047** (-2.24)	-0.047** (-2.20)	-0.035* (-1.95)	-0.043* (-1.98)	-0.052** (-2.30)	-0.053** (-2.25)	-0.050** (-2.20)	-0.052** (-2.19)	-0.041* (-1.73)	-0.053* (-1.93)
Number of observations	890	890	890	777	777	777	777	890	890	777
Adjusted R <sup>2</sup>	0.1728	0.1457	0.2161	0.1826	0.1499	0.1535	0.1643	0.1267	0.1337	0.1304

**Table 6. Determinants of global IPO activity: 1990 to 2007.**

The dependent variable is each country's annual measure of global IPO activity. IPO data is from SDC and includes 5,143 global IPOs from 54 countries that have data available for GDP and country  $q$  for at least one year during the sample period from 1990 to 2007. For each country, global IPO counts and proceeds are summed annually. Panel a shows regressions where the dependent variable is each country's annual global IPO count scaled by the total number of IPOs that year. Panel b shows regressions where the dependent variable is each country's annual global IPO proceeds scaled by the total number of IPO proceeds that year. Global IPO proceeds do not include proceeds from the domestic tranche of the IPO. Both measures of global IPO activity are multiplied by 100. The dependent variable is set to missing if there are no IPOs in a given country in a given year. In Panel a (Panel b), the world domestic IPO rate is based on counts (proceeds). The domestic IPO rate and world domestic IPO rate include total domestic proceeds in Panel b. With the exception of the institutions variables, all variables are lagged by one year. Variables are defined in Appendix B. The  $t$ -statistics (in parentheses) are adjusted for clustering on countries – they are computed assuming observations are independent across countries, but not within countries. \*, \*\*, and \*\*\* indicate statistical significance at the 10%, 5%, and 1% levels, respectively.

Panel a. Global IPO counts scaled by total number of IPOs.											
		Common law	Anti- director	Anti-self- dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	5.155 (0.27)	22.855 (1.18)	49.291** (2.20)	27.329 (1.23)	50.489** (2.24)	6.225 (0.30)	11.340 (0.45)	20.259 (1.02)	3.357 (0.18)	-25.837 (-1.00)	-82.465*** (-4.09)
Institutions variable		-17.477*** (-2.76)	-7.753*** (-2.81)	-28.999** (-2.67)	-53.909*** (-4.57)	-13.130 (-1.24)	-14.051 (-0.82)	-26.396** (-2.60)	0.162 (0.44)	-8.208* (-1.82)	101.053*** (5.97)
Domestic IPO rate	-3.336*** (-5.86)	-3.109*** (-5.79)	-3.370*** (-7.09)	-2.970*** (-5.13)	-3.354*** (-5.99)	-3.941*** (-6.74)	-3.915*** (-6.40)	-3.785*** (-6.52)	-3.383*** (-6.18)	-3.232*** (-5.59)	-3.485*** (-6.07)
World domestic IPO rate	1.948*** (2.98)	1.784*** (2.79)	1.979*** (3.13)	1.813*** (2.78)	1.246* (1.81)	1.529** (2.31)	1.592** (2.41)	1.504** (2.28)	1.996*** (3.06)	1.732** (2.57)	1.437** (2.07)
Country $q$	10.157* (1.73)	8.937 (1.42)	4.584 (0.81)	7.349 (1.13)	7.957 (1.31)	12.266** (2.03)	12.354* (1.83)	11.962* (1.85)	10.044 (1.66)	11.620* (2.00)	13.151** (2.14)
Market cap / GDP	-4.863 (-1.16)	0.892 (0.21)	0.780 (0.18)	0.423 (0.10)	6.093 (1.48)	-1.083 (-0.28)	-1.014 (-0.24)	1.712 (0.45)	-4.921 (-1.15)	-4.298 (-1.10)	-1.283 (-0.35)
Market turnover	-6.449** (-2.37)	-7.675*** (-2.76)	-6.391** (-2.36)	-6.884** (-2.60)	-6.127*** (-2.80)	-7.176** (-3.32)	-7.724*** (-3.25)	-7.609*** (-3.22)	-6.328** (-2.28)	-6.705** (-2.38)	-0.141 (-0.07)
Log (GDP / capita)	3.635* (1.79)	2.418 (1.27)	2.448 (1.33)	3.030 (1.42)	2.787 (1.43)	4.212** (2.21)	3.630* (1.82)	3.318* (1.78)	2.474 (0.69)	8.280** (2.51)	8.015*** (5.21)
World financial globalization	1.833 (0.77)	0.662 (0.30)	0.817 (0.34)	0.862 (0.38)	-1.389 (-0.61)	0.439 (0.19)	0.786 (0.34)	-0.064 (-0.03)	1.922 (0.80)	-0.906 (-0.34)	-2.145 (-0.77)
Number of observations	698	698	698	698	632	632	632	632	698	698	632
Adjusted R <sup>2</sup>	0.2160	0.2514	0.2519	0.2393	0.2917	0.2468	0.2461	0.2610	0.2156	0.2228	0.3269

**Table 6, continued.**

Panel b. Global IPO proceeds scaled by total IPO proceeds.											
		Common law	Anti-director	Anti-self-dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	8.481 (0.54)	21.261 (1.14)	31.188 (1.56)	24.112 (1.17)	50.711** (2.34)	14.264 (0.68)	9.488 (0.40)	23.590 (1.17)	11.673 (0.73)	-21.437 (-1.02)	-54.621*** (-2.91)
Institutions variable		-11.761** (-2.10)	-4.011* (-1.97)	-18.797** (-2.11)	-43.770*** (-4.21)	-11.216 (-1.10)	-3.332 (-0.23)	-18.995** (-2.20)	-0.328 (-0.95)	-7.998** (-2.30)	79.405*** (5.45)
Domestic IPO rate	-16.092*** (-3.92)	-14.235*** (-3.70)	-15.545*** (-4.17)	-13.469*** (-3.51)	-15.050*** (-4.04)	-17.725*** (-4.24)	-17.828*** (-3.92)	-16.847*** (-4.07)	-15.328*** (-3.75)	-15.389*** (-3.82)	-17.085*** (-4.09)
World domestic IPO rate	48.917*** (3.01)	43.727** (2.65)	48.011*** (2.90)	44.931** (2.67)	35.227** (2.03)	42.915** (2.51)	44.232** (2.63)	40.611** (2.41)	50.164*** (2.97)	44.095*** (2.78)	37.618** (2.29)
Country $q$	0.371 (0.07)	-0.353 (-0.06)	-2.656 (-0.48)	-1.245 (-0.21)	-1.673 (-0.31)	0.967 (0.17)	1.259 (0.21)	1.014 (0.17)	0.619 (0.11)	1.913 (0.36)	2.615 (0.52)
Market cap / GDP	0.285 (0.07)	4.010 (0.93)	3.055 (0.70)	3.523 (0.81)	8.914** (2.25)	2.899 (0.71)	1.742 (0.41)	4.650 (1.12)	0.403 (0.11)	0.883 (0.23)	3.094 (1.05)
Market turnover	-10.355*** (-6.46)	-11.110*** (-6.40)	-10.381*** (-6.25)	-10.507*** (-6.59)	-8.928*** (-5.21)	-9.946*** (-6.27)	-10.288*** (-6.05)	-10.293*** (-6.28)	-10.484*** (-6.64)	-10.503*** (-6.19)	-4.301** (-2.41)
Log (GDP / capita)	2.482 (1.62)	1.640 (1.03)	1.920 (1.28)	2.007 (1.24)	1.420 (0.85)	2.635 (1.59)	2.657 (1.57)	2.019 (1.21)	4.774 (1.64)	6.959*** (2.77)	5.581*** (4.14)
World financial globalization	5.068** (2.47)	4.368** (2.15)	4.574** (2.18)	4.457** (2.14)	1.794 (0.85)	3.175 (1.52)	3.620* (1.74)	2.865 (1.37)	5.035** (2.48)	2.646 (1.16)	0.917 (0.43)
Number of observations	700	700	700	700	633	633	633	633	700	700	633
Adjusted R <sup>2</sup>	0.1033	0.1234	0.1150	0.1155	0.1521	0.1104	0.1044	0.1184	0.1061	0.1119	0.1774

**Table 7. Determinants of global IPO activity and financial globalization.**

The dependent variable is each country's annual measure of global IPO activity. IPO data is from SDC and includes 5,143 global IPOs from 54 countries that have data available for GDP and for country  $q$  for at least one year during the sample period from 1990 to 2007. For each country, global IPO counts and proceeds are summed annually. Panel a shows regressions where the dependent variable is each country's annual global IPO count scaled by the total number of IPOs that year. Panel b shows regressions where the dependent variable is each country's annual global IPO proceeds scaled by the total number of IPO proceeds that year. Global IPO proceeds do not include proceeds from the domestic tranche of the IPO. Both measures of global IPO activity are multiplied by 100. The dependent variable is set to missing if there are no IPOs in a given country in a given year. In Panel a (Panel b), the world domestic IPO rate is based on counts (proceeds). The domestic IPO rate and world domestic IPO rate include total domestic proceeds in Panel b. With the exception of the institutions variables, all variables are lagged by one year. Variables are defined in Appendix B. The  $t$ -statistics (in parentheses) are adjusted for clustering on countries – they are computed assuming observations are independent across countries, but not within countries. \*, \*\*, and \*\*\* indicate statistical significance at the 10%, 5%, and 1% levels, respectively.

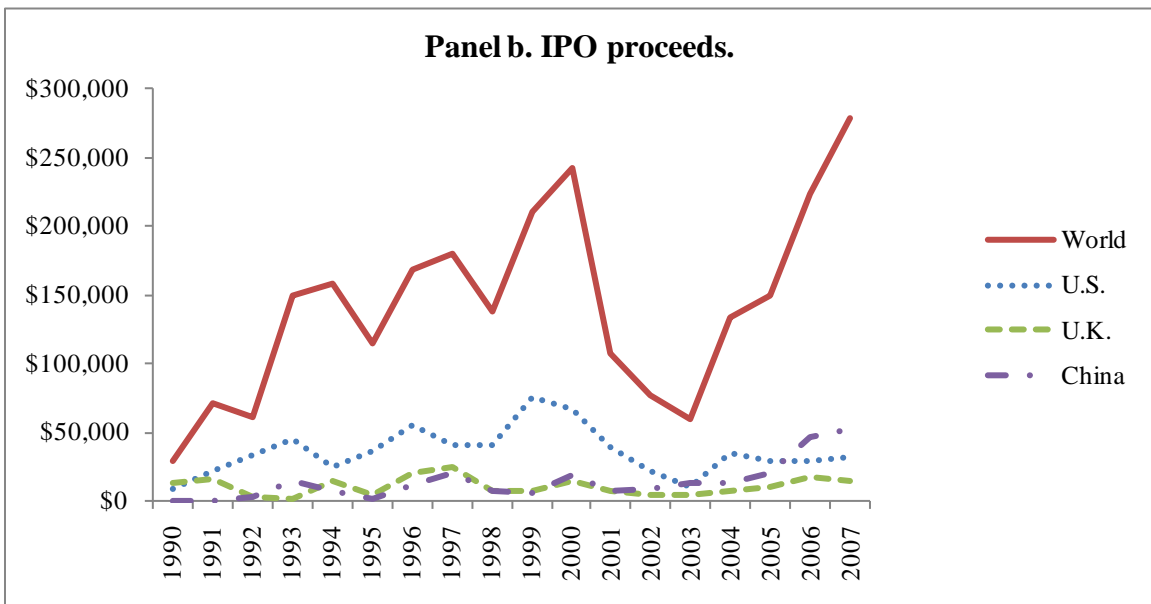
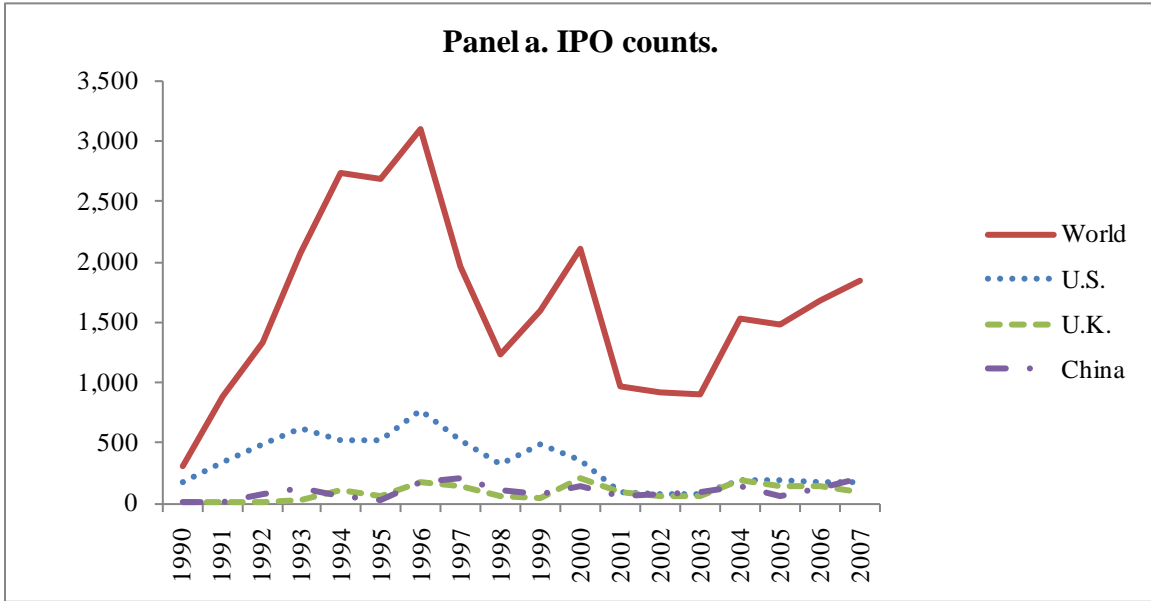
Panel a. Global IPO counts scaled by total number of IPOs.										
	Common law	Anti-director	Anti-self-dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	18.551 (1.04)	23.462 (1.33)	13.863 (0.69)	13.824 (0.78)	0.739 (0.04)	5.662 (0.29)	7.279 (0.41)	18.705 (0.58)	-33.944 (-1.13)	-39.331** (-2.33)
Institutions variable	-17.980*** (-2.86)	-7.976*** (-2.88)	-29.623*** (-2.76)	-53.876*** (-4.58)	-13.077 (-1.23)	-14.013 (-0.82)	-26.458** (-2.58)	0.162 (0.44)	-8.380* (-1.82)	100.828*** (6.01)
Institutions × World financial globalization	9.325** (2.21)	4.984** (2.47)	14.063** (2.14)	3.702 (0.44)	-4.167 (-0.56)	-1.420 (-0.17)	2.808 (0.34)	-0.025 (-0.15)	0.987 (0.44)	-13.467 (-1.22)
Domestic IPO rate	-3.038*** (-5.78)	-3.368*** (-7.23)	-2.939*** (-5.08)	-3.338*** (-6.03)	-3.952*** (-6.75)	-3.920*** (-6.35)	-3.771*** (-6.46)	-3.379*** (-6.14)	-3.243*** (-5.62)	-3.504*** (-6.09)
World domestic IPO rate	1.714*** (2.68)	1.879*** (3.00)	1.779*** (2.75)	1.225* (1.81)	1.548** (2.35)	1.596** (2.44)	1.491** (2.28)	1.992*** (3.02)	1.727** (2.57)	1.416** (2.05)
Country $q$	9.472 (1.53)	5.632 (0.99)	7.945 (1.23)	8.181 (1.32)	12.115* (1.97)	12.277* (1.82)	12.099* (1.86)	10.101 (1.67)	11.500* (1.99)	13.190** (2.18)
Market cap / GDP	0.635 (0.16)	0.741 (0.18)	0.257 (0.06)	5.977 (1.47)	-0.990 (-0.26)	-0.983 (-0.23)	1.630 (0.43)	-4.909 (-1.14)	-4.383 (-1.12)	-1.323 (-0.36)
Market turnover	-8.159*** (-3.03)	-6.851** (-2.57)	-6.869** (-2.58)	-6.167*** (-2.85)	-7.120*** (-3.29)	-7.715*** (-3.26)	-7.660*** (-3.27)	-6.347** (-2.27)	-6.733** (-2.40)	-0.328 (-0.16)
Log (GDP / capita)	2.435 (1.30)	2.456 (1.36)	3.076 (1.45)	2.811 (1.45)	4.208** (2.20)	3.626* (1.82)	3.327* (1.79)	2.492 (0.69)	8.334** (2.53)	7.985*** (5.20)
World financial globalization	0.509 (0.23)	0.029 (0.01)	0.442 (0.20)	-1.483 (-0.65)	0.536 (0.23)	0.791 (0.35)	-0.110 (-0.05)	1.957 (0.82)	-1.058 (-0.40)	-2.342 (-0.83)
Number of observations	698	698	698	632	632	632	632	698	698	632
Adjusted R <sup>2</sup>	0.2557	0.2588	0.2415	0.2907	0.2459	0.2449	0.2599	0.2145	0.2219	0.3268

**Table 7, continued.**

Panel b. Global IPO proceeds scaled by total IPO proceeds.										
	Common law	Anti-director	Anti-self-dealing	Disclosure	Burden of proof	Public enforce	Investor protection	Political risk	Rule of law	Ownership
Constant	25.857 (1.52)	25.616 (1.57)	22.822 (1.30)	25.882 (1.51)	14.776 (0.87)	14.460 (0.78)	19.535 (1.14)	-2.769 (-0.10)	-22.538 (-0.92)	-15.551 (-1.00)
Institutions variable	-12.207** (-2.21)	-4.107** (-2.04)	-19.377** (-2.20)	-43.570*** (-4.18)	-11.263 (-1.11)	-3.418 (-0.24)	-19.126** (-2.22)	-0.328 (-0.95)	-8.003** (-2.27)	79.324*** (5.42)
Institutions × World financial globalization	8.491** (2.30)	2.487 (1.43)	10.565* (1.67)	9.008 (1.11)	4.201 (0.62)	2.561 (0.37)	7.068 (0.96)	0.028 (0.15)	0.026 (0.01)	-16.025* (-1.72)
Domestic IPO rate	-13.312*** (-3.49)	-15.278*** (-4.16)	-12.768*** (-3.24)	-14.623*** (-3.92)	-17.560*** (-4.17)	-17.664*** (-3.83)	-16.442*** (-3.93)	-15.341*** (-3.76)	-15.388*** (-3.82)	-17.199*** (-4.08)
World domestic IPO rate	42.669** (2.61)	47.039*** (2.85)	44.175** (2.62)	34.336* (2.00)	42.478** (2.51)	44.049** (2.63)	39.944** (2.40)	50.338*** (2.97)	44.083*** (2.79)	36.582** (2.22)
Country $q$	0.106 (0.02)	-2.198 (-0.40)	-0.861 (-0.14)	-1.108 (-0.20)	1.120 (0.19)	1.396 (0.23)	1.374 (0.23)	0.545 (0.10)	1.910 (0.35)	2.646 (0.52)
Market cap / GDP	3.685 (0.88)	2.979 (0.70)	3.311 (0.78)	8.591** (2.15)	2.791 (0.68)	1.669 (0.39)	4.415 (1.07)	0.385 (0.10)	0.881 (0.23)	3.042 (1.01)
Market turnover	-11.508*** (-7.03)	-10.612*** (-6.46)	-10.501*** (-6.50)	-9.012*** (-5.31)	-9.999*** (-6.34)	-10.302*** (-6.03)	-10.408*** (-6.49)	-10.467*** (-6.70)	-10.504*** (-6.12)	-4.522** (-2.57)
Log (GDP / capita)	1.647 (1.05)	1.944 (1.30)	2.039 (1.26)	1.479 (0.88)	2.639 (1.59)	2.663 (1.57)	2.040 (1.22)	4.762 (1.63)	6.961*** (2.78)	5.558*** (4.12)
World financial globalization	4.336** (2.18)	4.289** (2.08)	4.202** (2.08)	1.661 (0.79)	3.113 (1.48)	3.628* (1.73)	2.810 (1.35)	4.994** (2.40)	2.642 (1.14)	0.721 (0.34)
Number of observations	700	700	700	633	633	633	633	700	700	633
Adjusted R <sup>2</sup>	0.1280	0.1163	0.1166	0.1520	0.1094	0.1031	0.1181	0.1048	0.1106	0.1779

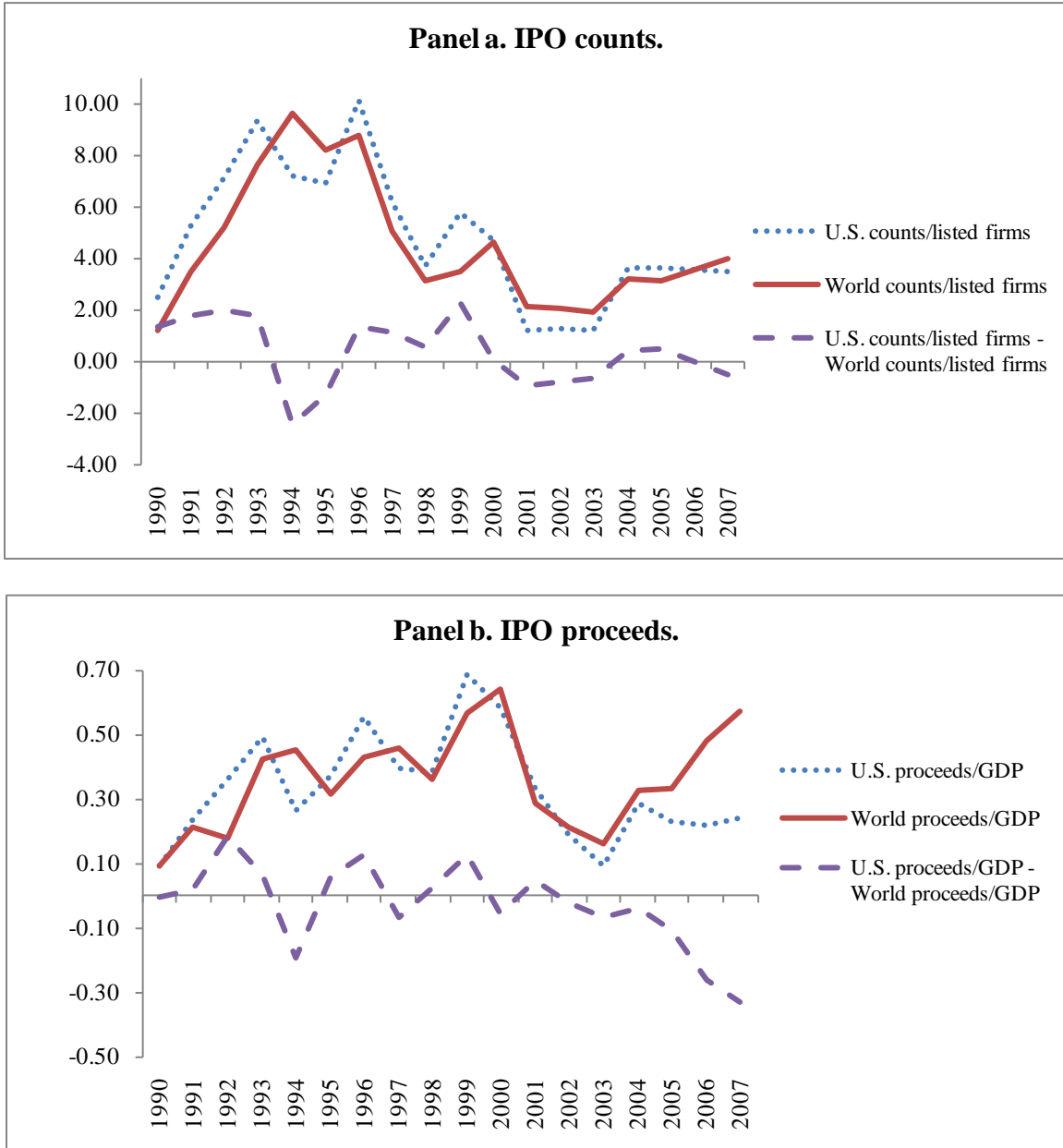
**Figure 1. Total IPO activity: 1990 to 2007.**

This figure shows annual IPO activity for all countries (World), the U.S., U.K., and China from 1990 to 2007. IPO data is from SDC and includes 29,361 IPOs from 89 countries over the period from 1990 to 2007. Panel a shows the total number of IPOs (domestic and global) each year. Panel b shows total IPO proceeds raised (domestic and global) each year. Proceeds are in constant 2007 U.S. dollars (millions).



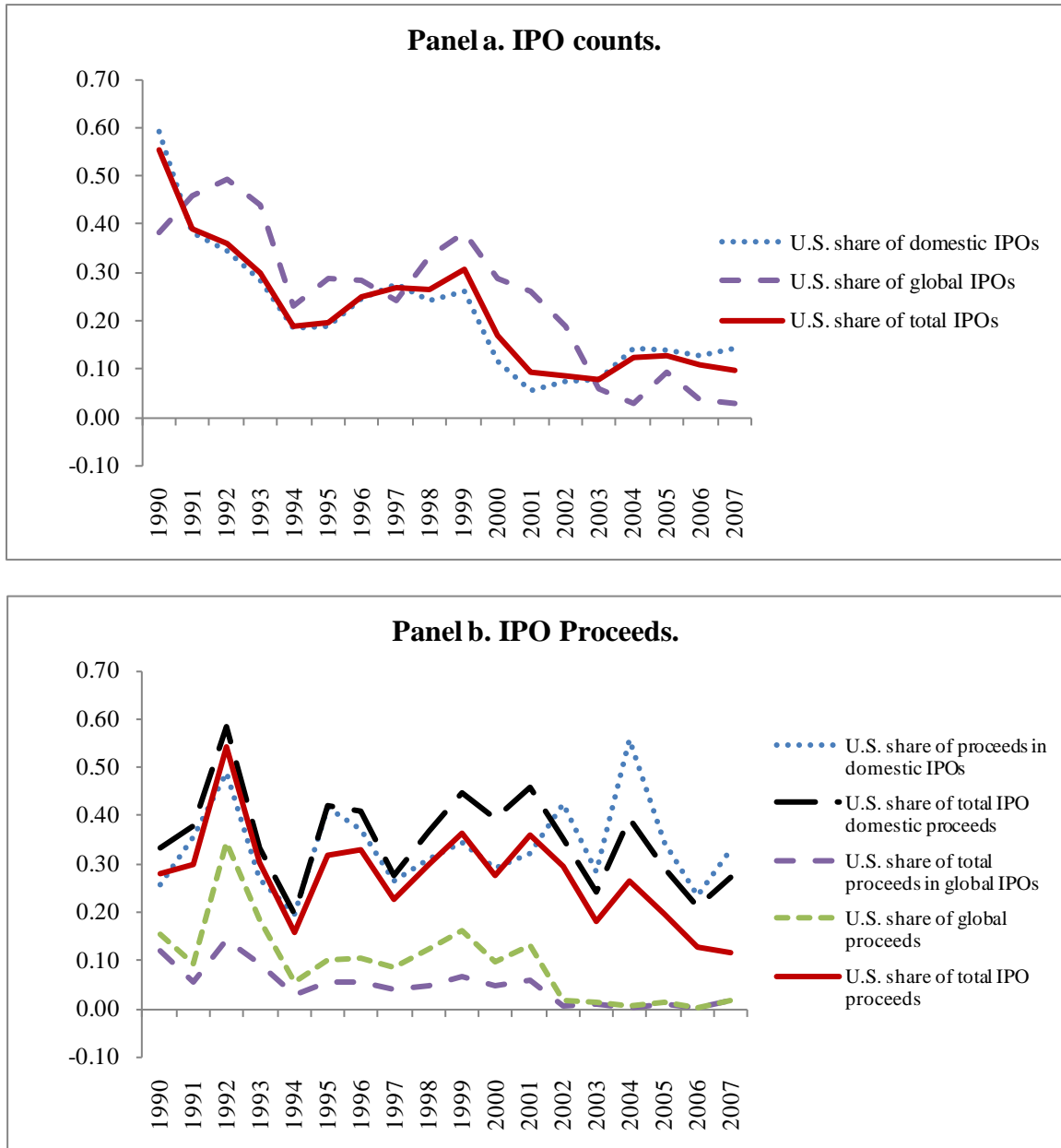
**Figure 2. U.S. and world IPO activity: 1990 to 2007.**

This figure shows annual IPO activity for the U.S. and all countries (World) from 1990 to 2007. IPO data is from SDC and includes 29,361 IPOs from 89 countries. Panel a shows the total number of IPOs (domestic and global) scaled by the lagged number of domestic firms each year. Panel b shows total IPO proceeds raised (domestic and global) scaled by lagged GDP each year. Proceeds are in constant 2007 U.S. dollars (millions).



**Figure 3. The U.S. share of world IPO activity.**

This figure shows annual share of IPO activity for U.S. firms relative to firms in the rest of the world from 1990 to 2007. IPO data is from SDC and includes 29,361 IPOs from 89 countries. Panel a shows the number of U.S. IPOs scaled by the number of IPOs by firms from the rest of the world. Panel b shows IPO proceeds raised by U.S. firms scaled by IPO proceeds raised by firms from the rest of the world. Proceeds and GDP are in constant 2007 U.S. dollars (millions).



**Appendix A. Summary statistics and correlations for country-level variables.**

This table shows the average values of the country variables. The sample is restricted to 54 countries that have data available for GDP and for country  $q$  for at least one year during the sample period from 1990 to 2007. Each variable is averaged across years within a given country and is then averaged across countries.

Variable	N	Mean	Median	Std dev	Min	Max	25 <sup>th</sup> pctile	75 <sup>th</sup> pctile
Domestic IPO counts scaled by the lagged number of domestic firms	54	2.567	1.317	2.879	0.019	11.959	0.366	3.929
Domestic IPO proceeds scaled by lagged GDP	54	0.140	0.094	0.142	0.000	0.628	0.040	0.174
Global IPO counts scaled by total number of IPOs	54	47.924	49.284	26.626	1.678	92.941	25.000	69.608
Global IPO proceeds scaled by total IPO proceeds	54	45.702	44.591	20.764	6.487	94.077	32.816	57.460
World domestic IPO rate (counts)	54	3.686	3.735	0.225	2.353	3.944	3.710	3.750
World domestic IPO rate (proceeds)	54	0.153	0.153	0.003	0.142	0.161	0.153	0.154
World domestic IPO rate (total domestic proceeds)	54	0.244	0.244	0.004	0.222	0.255	0.244	0.245
Domestic IPO rate (counts)	54	2.490	1.169	2.849	0.000	12.271	0.347	3.879
Domestic IPO rate (total domestic proceeds)	54	0.205	0.152	0.179	0.004	0.890	0.070	0.265
Common law	53	0.302	0.000	0.463	0.000	1.000	0.000	1.000
Anti-director	53	3.443	3.500	1.121	1.000	5.000	3.000	4.000
Anti-self-dealing	53	0.484	0.440	0.240	0.092	1.000	0.288	0.642
Disclosure	45	0.624	0.580	0.209	0.170	1.000	0.500	0.750
Burden of proof	45	0.487	0.440	0.252	0.000	1.000	0.220	0.660

**Appendix A, continued.**

Variable	N	Mean	Median	Std dev	Min	Max	25 <sup>th</sup> pctile	75 <sup>th</sup> pctile
Public enforcement	45	0.517	0.550	0.223	0.000	0.900	0.333	0.667
Investor protection	45	0.481	0.465	0.234	0.000	1.000	0.355	0.610
Political risk	53	73.498	75.685	11.616	47.468	92.382	66.067	83.519
Rule of law	54	0.763	0.856	0.925	-1.040	1.986	-0.010	1.643
Ownership	45	0.462	0.510	0.132	0.180	0.670	0.390	0.560
Country $q$	54	1.279	1.294	0.189	0.884	1.847	1.177	1.362
Market cap / GDP	54	0.594	0.429	0.515	0.072	2.587	0.220	0.815
Market turnover	54	0.588	0.482	0.494	0.016	2.745	0.234	0.722
Log (GDP / capita)	54	8.884	9.284	1.377	6.082	10.782	7.967	10.134
World financial globalization	54	1.893	1.893	0	1.893	1.893	1.893	1.893

**Appendix A, continued.**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)	(17)	(18)	(19)	(20)	(21)	(22)	(23)	(24)
Dom IPO counts / lagged # of dom firms (1)	1.00																							
Dom IPO proceeds / lagged GDP (2)	0.79	1.00																						
Global IPO counts / total # if IPOs (3)	-0.75	-0.67	1.00																					
Global IPO proceeds / total IPO proceeds (4)	-0.56	-0.50	0.86	1.00																				
World dom IPO rate (counts) (5)	-0.40	-0.21	0.24	0.10	1.00																			
World dom IPO rate (proceeds) (6)	-0.33	-0.50	0.43	0.39	0.19	1.00																		
World dom IPO rate (total dom proceeds) (7)	-0.23	-0.34	0.27	0.30	0.11	0.84	1.00																	
Dom IPO rate (counts) (8)	1.00	0.79	-0.75	-0.56	-0.41	-0.34	-0.24	1.00																
Dom IPO rate (total dom proceeds) (9)	0.71	0.92	-0.56	-0.43	-0.18	-0.41	-0.38	0.71	1.00															
Common law (10)	0.38	0.48	-0.41	-0.32	-0.38	-0.46	-0.41	0.39	0.44	1.00														
Anti-director (11)	0.34	0.39	-0.34	-0.20	-0.27	-0.15	-0.02	0.35	0.39	0.54	1.00													
Anti-self-dealing (12)	0.50	0.60	-0.30	-0.16	-0.28	-0.38	-0.34	0.51	0.61	0.72	0.64	1.00												
Disclosure (13)	0.62	0.62	-0.50	-0.45	-0.35	-0.41	-0.42	0.62	0.62	0.64	0.50	0.64	1.00											
Burden of proof (14)	0.38	0.37	-0.23	-0.17	-0.20	-0.55	-0.45	0.38	0.40	0.34	0.39	0.37	0.49	1.00										
Public enforcement (15)	0.35	0.39	-0.23	-0.08	-0.14	-0.35	-0.51	0.35	0.43	0.40	0.13	0.37	0.44	0.34	1.00									
Investor protection (16)	0.46	0.47	-0.34	-0.23	-0.28	-0.57	-0.56	0.46	0.50	0.58	0.50	0.58	0.62	0.79	0.71	1.00								
Political risk (17)	0.19	0.14	0.10	0.10	-0.07	0.01	-0.08	0.17	0.30	-0.13	0.00	0.14	0.13	0.17	-0.23	-0.01	1.00							
Rule of law (18)	0.24	0.18	0.04	0.04	-0.12	-0.02	-0.13	0.23	0.34	0.04	0.13	0.27	0.26	0.23	-0.16	0.11	0.92	1.00						
Ownership (19)	-0.44	-0.28	0.44	0.46	0.26	0.37	0.34	-0.42	-0.33	-0.16	-0.24	-0.25	-0.43	-0.38	-0.04	-0.36	-0.47	-0.51	1.00					
Country <i>q</i> (20)	0.09	-0.04	-0.01	-0.08	-0.18	-0.15	-0.19	0.10	-0.06	0.11	-0.07	0.00	0.17	0.10	0.04	0.12	-0.07	0.07	-0.09	1.00				
Market cap / GDP (21)	0.36	0.53	-0.14	-0.03	-0.03	-0.28	-0.28	0.37	0.59	0.36	0.35	0.52	0.59	0.40	0.26	0.46	0.38	0.41	-0.30	-0.01	1.00			
Market turnover (22)	0.43	0.31	-0.30	-0.38	-0.13	-0.14	-0.13	0.42	0.22	-0.05	0.09	0.01	0.21	0.16	-0.04	0.08	0.12	0.18	-0.57	0.21	0.05	1.00		
Log (GDP / capita) (23)	0.13	0.06	0.18	0.22	0.04	0.05	-0.08	0.11	0.22	-0.14	-0.03	0.15	0.06	0.13	-0.21	-0.03	0.91	0.87	-0.44	-0.11	0.37	0.17	1.00	
World financial globalization (24)	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.	.

## Appendix B. Variable definitions.

IPO data is from SDC's Global New Issues Database. IPO proceeds are in constant 2007 U.S. dollars. Country-level variables are from the World Bank's WDI Database, La Porta, Lopez-de-Silanes, and Shleifer (2006), Djankov, La Porta, Lopez-de-Silanes, and Shleifer (2008), International Country Risk Guide database, the World Bank's World Governance Indicators database, and the 2008 update of the Financial Development and Structure database, originally used in Beck, Demirgüç-Kunt, and Levine (2000). Data to compute Tobin's  $q$  is from Worldscope.

Variable	Definition
Domestic IPO counts / lagged # of domestic firms	Number of domestic IPOs in country $j$ in year $t$ / number of domestic listed firms in country $j$ in year $t-1$ and is multiplied by 100. Source: SDC and WDI database.
Domestic IPO proceeds / lagged GDP	Proceeds raised in domestic IPOs in country $j$ in year $t$ / GDP for country $j$ in year $t-1$ and is multiplied by 100. Source: SDC and WDI database.
Global IPO counts / total number of IPOs	Number of global IPOs in country $j$ in year $t$ / the total number of IPOs in country $j$ in year $t$ and is multiplied by 100. Source: SDC.
Global IPO proceeds / total IPO proceeds	Global IPO proceeds raised in country $j$ in year $t$ / total IPO proceeds raised in country $j$ in year $t$ and is multiplied by 100. Global IPO proceeds include proceeds raised in the international tranches only. Source: SDC.
World domestic IPO rate (counts)	Total world domestic IPO counts in year $t$ / total number of domestic listed firms worldwide in year $t-1$ and is multiplied by 100. To compute the world domestic IPO rate for country $j$ , domestic IPOs and the number of domestic listed firms for country $j$ are excluded from the calculation. Used in Tables 4 and 5. Source: SDC and WDI database.
World domestic IPO rate (proceeds)	Total world proceeds raised in domestic IPOs in year $t$ / total worldwide GDP in year $t-1$ and is multiplied by 100. To compute the world domestic IPO rate for country $j$ , IPO proceeds and GDP for country $j$ are excluded from the calculation. Used in Tables 4 and 5. Source: SDC and WDI database.
World domestic IPO rate (total domestic proceeds)	Total world domestic IPO proceeds in year $t$ / total worldwide GDP in year $t-1$ and is multiplied by 100. Total world domestic IPO proceeds include proceeds raised in domestic IPOs and the domestic component of global IPOs. To compute the world domestic IPO rate for country $j$ , IPO proceeds and GDP for country $j$ are excluded from the calculation. Used in Tables 6 and 7. Source: SDC and WDI database.
Domestic IPO rate (counts)	Lagged domestic IPO counts / lagged # of domestic firms and is multiplied by 100. Source: SDC and WDI database.
Domestic IPO rate (total domestic proceeds)	Lagged domestic IPO proceeds / lagged GDP and is multiplied by 100. For this variable, proceeds include total domestic proceeds, including proceeds raised in domestic IPOs and the domestic component of global IPOs. Source: SDC and WDI database.
U.S. dummy 1990s	Equals one for the US for 1990 to 1999 and zero otherwise.
U.S. dummy 2000s	Equals one for the US for 2000 to 2007 and zero otherwise.
Common law	Equals one if a country's origin of commercial law is English common law, and zero otherwise. Source: DLLS (2008).
Anti-director	The index is formed by summing: (1) vote by mail; (2) shares not deposited; (3) cumulative voting; (4) oppressed minority; (5) pre-emptive rights; and (6) capital to call a meeting. Ranges from zero to six. Source: DLLS (2008).
Anti-self-dealing	Average of ex ante and ex post private control of self-dealing, where ex ante is the average of approval by disinterested shareholders and ex ante disclosure; ex post is the average of disclosure in periodic filings and ease of proving wrongdoing. Ranges from zero to one. Source: DLLS (2008).

**Appendix B, continued.**

Variable	Definition
Anti-self-dealing	Average of ex ante and ex post private control of self-dealing, where ex ante is the average of approval by disinterested shareholders and ex ante disclosure; ex post is the average of disclosure in periodic filings and ease of proving wrongdoing. Ranges from zero to one. Source: DLLS (2008).
Disclosure	Arithmetic mean of (1) prospectus; (2) compensation; (3) shareholders; (4) inside ownership; (5) contracts irregular; and (6) transactions. Ranges from zero to one. Source: LLS (2006).
Burden of proof	Arithmetic mean of (1) liability standard for the issuer and its directors; (2) liability standard for distributors; and (3) liability standard for accountants. Ranges from zero to one. Source: LLS (2006).
Public enforcement	Arithmetic mean of (1) supervisor characteristics index; (2) rule-making power index; (3) investigative powers index; (4) orders index; and (5) criminal index. Ranges from zero to one. Source: LLS (2006).
Investor protection	Principal component of disclosure, burden of proof, and anti-director rights. Ranges from zero to one. Source: LLS (2006).
Political risk	Includes 12 weighted variables covering both political and social attributes. Ranges from zero to 100. Source: ICRG.
Rule of law	Captures perceptions of the extent to which agents have confidence in and abide by the rules of society, and in particular the quality of contract enforcement, property rights, the police, and the courts, as well as the likelihood of crime and violence. Ranges from -1.6753 to 2.0431. Source: 2009 update of the Worldwide Governance Indicators database.
Ownership	Average percentage of common shares owned by the top three shareholders in the 10 largest nonfinancial, privately owned domestic firms in a given country. Source: LLSV (1998).
Country $q$	For each firm in country $j$ $q$ is computed annually as total assets less the book value of equity plus the market value of equity / book value of total assets (all variables in local currency). For each country, median industry $qs$ are computed annually using the Fama-French 17 industry classification scheme. The industry $qs$ are then weighted by their relative market values each year so that country $q$ is the market value weighted average of the median industry $qs$ . Source: Worldscope.
Market cap / GDP	Value of listed shares to GDP. Source: Financial Development and Structure database.
Market turnover	Ratio of the value of total shares traded to average real market capitalization. Source: Financial Development and Structure database.
Log (GDP / capita)	Gross domestic product divided by midyear population. GDP is in current U.S. dollars. Source: WDI database.
World financial globalization	Each year, the U.S. dollar-denominated value of external assets and liabilities is summed across countries and divided by world GDP (in U.S. dollars). Source: updated and extended dataset compiled by Lane and Milesi-Ferretti (2007).